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METHODS FOR DETERMINING TIME TO RETURN TO PLAY AFTER RECREATIONAL INJURY IN FIELD AND COURT SPORT ATHLETES

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METHODS FOR DETERMINING TIME TO RETURN TO PLAY AFTER RECREATIONAL INJURY IN FIELD AND COURT SPORT ATHLETES

_____________________________________
DISSERTATION
_____________________________________

A dissertation submitted in partial fulfillment of the requirements for the degree of Doctor of Philosophy in the College of Public Health at the University of Kentucky

By Sarah Nicole Morris

Lexington, Kentucky

Co-Directors: Dr. Wayne T. Sanderson, Professor of Epidemiology and Dr. Mary Kay Rayens, Professor of Biostatistics

Lexington, Kentucky

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ABSTRACT OF DISSERTATION

METHODS FOR DETERMINING TIME TO RETURN TO PLAY AFTER RECREATIONAL INJURY IN FIELD AND COURT SPORT ATHLETES

An observational study was used to illustrate the application of time to event analysis methods to return to play; a secondary data analysis of athlete injury data from the High School RIO™ Injury Surveillance System (ISS) database was conducted. National Athletic Trainers’ Association (NATA)-certified athletic trainers from approximately 100 high schools in the US enroll their school in the system and complete the online “Exposure Report Form” for reportable injuries each week. New lateral ankle sprains and single-ligament knee injuries experienced by high school athletes during regularly scheduled participation in school-sanctioned sports for seven academic years (2005-2006 through 2011-2012) were analyzed. Field and court sport athletes (football, boys/girls soccer, volleyball, wrestling, basketball, baseball, and softball) were considered as these athletes were more likely to suffer lateral ankle or knee ligament sprains.

Detailed guidance was provided to assist athletic trainers and sports medicine researchers with understanding the appropriate data structure and programming statements required for time to return to play (T-RTP) analysis and the methodology appropriate for analyzing discrete time RTP categories. A data example was presented using lateral ankle sprain information to demonstrate how the life-table is useful for generating directly applicable information on expected T-RTP, and a discrete logistic regression model for this example highlights the relationship between severity of injury and T-RTP. Coding statements and life-table output were detailed for the LIFETEST procedure in SAS; SPSS instructions for generating life-tables were documented. The PHREG procedure in SAS using the TIES=DISCRETE option was presented to generate the discrete logistic regression model. An alternative method for computing hazard odds ratios was discussed to reduce computing time for large datasets with high numbers of tied event times using a pseudo dataset and the LOGISTIC procedure.

For 1st and 2nd degree lateral ankle sprains, the probability of RTP was highest 10-21 days after injury. For 3rd degree lateral ankle sprain, the probability of RTP was highest at least four weeks after injury. Gender had a marginal effect on RTP; male athletes were 18% more likely to return to play than female athletes. There was a significant interaction effect on RTP between time interval of return and ankle sprain severity. Athletes who
experienced a 1st degree sprain were 458% more likely to RTP in 1-2 days than athletes who experienced a 3rd degree sprain, and 2nd degree sprains were 259% more likely to RTP in 1-2 days than 3rd degree sprains. In general, 1st and 2nd degree LAS were more likely to return than 3rd degree sprains in the three weeks after injury.

Regardless of which knee ligament was injured, athletes had a very small chance of RTP within two weeks of injury. Athletes injuring the ACL any time during the season had only a 1 in 3 chance of returning before the end of the season. RTP probabilities increase slightly for PCL, LCL, and MCL injuries after two weeks. Athletes suffering a single-ligament knee sprain during competition were 25% less likely to RTP before the end of the season than athletes injured during practice. Gender did not have a significant effect on RTP. There was a significant interaction effect on RTP between time interval of return and injured knee ligament. Athletes who experienced ACL sprain were 78% less likely to RTP in 1-2 days than athletes with MCL sprain, 81% less likely to return in 3-6 days, 91% less likely to return in 7-21 days, and 74% less likely to return 4 weeks after injury. Athletes who experienced LCL sprains were 213% more likely to return in 1-2 days than athletes with MCL sprain, 73% more likely to return in 3-6 days, and 103% more likely to return in 7-9 days.

The literature on return to play has been largely descriptive in nature, and time to event analysis methodology has not been heavily utilized. The applied methods paper presented here provides sports medicine researchers with direction to apply the methodology and interpret the results. The findings suggest that ankle sprain severity has the strongest impact on RTP timelines. ACL sprains have the longest RTP times and athletes are not likely to return during the season; athletes who suffer MCL sprains could potentially return during the season, but can expect to be out a minimum of three weeks. These RTP probability estimates are directly applicable for use by coaches, athletic trainers, and other members of the sports medicine team as they help provide reasonable expectations for return time following injury and allow for more accurate RTP planning.

KEYWORDS: sports medicine, knee sprain, ankle sprain, time to return to play, life-table method, discrete logistic regression

Sarah Nicole Morris
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April 14, 2017
Date
METHODS FOR DETERMINING TIME TO RETURN TO PLAY AFTER RECREATIONAL INJURY IN FIELD AND COURT SPORT ATHLETES

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April 14, 2017
This dissertation is dedicated to my mom, Gina, a brilliant example of what a woman can accomplish; and to my dad, Mike, who would have enjoyed seeing this goal finally achieved. Thank you for teaching me that I can do anything I set my mind to.
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1 Introduction

Ankle and knee injuries are ubiquitous with sports participation and account for up to 60% of all injuries that occur during play [1-3]. Ankle and knee sprains are the most common injuries among athletes of all ages [4-7] across all sports [1]. An initial question after an athlete sustains an injury is when they will be cleared to return to play (RTP). The extent of tissue damage and tissue healing timelines are not direct indicators of how long an athlete will be withheld from play, and many RTP decisions focus on resolving the symptoms. However, symptom and function resolution can follow a vastly different timeline than tissue healing, making a prognosis of time to RTP difficult. Current clinical predictions of when an athlete will return to play after ankle and knee injuries are likely to be misleading and erroneously underestimate RTP timelines in instances where a lack of follow-up data excludes athletes from analysis. Current RTP estimations are predominantly based on subjective reports of the individual patient’s symptoms and anecdotal, expert opinion of clinicians. While neither of these approaches are incorrect, there is limited clinical epidemiological research evidence to substantiate current RTP predictions for ankle and knee injuries [8]. There are few prognostic indicators for these injuries for determining when RTP will occur. This can potentially contribute to poor compliance with rehabilitation programs due to athletes pushing for earlier RTP, setting the athlete up for subsequent injury or increasingly severe injuries. Recurrent knee and ankle injuries can lead to concomitant long-term health issues such as degenerative joint diseases like osteoarthritis (OA). Ligament damage to the ankle or knee will likely result in early-onset, post-traumatic OA, most often within 10 to 20 years after initial injury [9, 10]. For high school or college athletes suffering ankle or knee injuries, this indicates
they could suffer degenerative conditions by as early as 30 years old. While OA is generally considered a factor of old age characterized by joint pain, swelling, limited motion, and disability [11], young athletes are at risk for early-onset consequences associated with ankle and knee joint injuries. Individuals suffering from posttraumatic OA are typically younger patients, and the condition is significantly associated with decreased physical activity, obesity, cardiovascular disease, and depression stemming from loss of function and disability due to injury to the affected joint [12, 13]. With approximately 30 million youth between the ages of 5 and 18 years in the US participating in organized sports [7], the potential is quite high for young athletes to suffer acute sports-related injuries or overuse injuries. It is estimated that 38% of high school athletes will suffer sports-related injuries requiring treatment by a physician [7]; however, actual percentages may be higher due to underreporting or failure to seek treatment [14].

To address these potentially significant public health issues related to knee and ankle sprains among high school athletes, there is a need for the development of effective strategies to diminish the impact of joint injuries, improve compliance with rehabilitation programs, and reduce the risk of reinjury, with the goal of avoiding long-term effects from injury and the continued maintenance of joint health. It is critical that athletic trainers, doctors, and coaches be able to accurately gauge when an athlete should return to play.

The purpose of this dissertation is to review the existing literature related to RTP after ankle and knee sprain, identify factors that affect RTP timelines, and generate evidence-based, objective prognostic indicators of when an athlete is likely to return to
participation using time to event analysis methodologies. The specific aims of this study are:

1. Examine return to play and the use of time to event analysis methodology in existing athletic training and sports medicine literature.
2. Provide guidance on generating and reporting return to play probabilities using time to event analysis methodology for athletic trainers and sports medicine researchers.
3. Analyze return to play probabilities for lateral ankle sprain and single-ligament knee injuries in high school athletes participating in field and court sports.

1.1 Motivation for dissertation

Return to play has historically been determined using subjective reasoning; there is a need for more objective methods to assist in the determination of RTP [8]. Typical research studies involving RTP report rates or proportions, but these measures can be inaccurate in instances where a lack of follow-up data excludes some athletes from analysis. Time to event analysis is a commonly used statistical analysis method that can provide more accurate estimates for time to return to play (T-RTP) by accounting for all injured athletes regardless of lack of follow-up concerning RTP. These analysis techniques have commonly not been applied or related to RTP, and will provide athletic trainers, coaches, and team physicians a more accurate way to estimate return.

Traditionally, time to event analyses have been aimed at estimating probabilities of negative events such as death, recurrence of illness, or recidivism. In the sports medicine
setting the return to play event is positive, requiring a shift in the interpretation of time to event analysis results. These analysis techniques will be used to conduct a secondary analysis on data from the High School RIO (Reporting Information Online)™ Injury Surveillance System (ISS) database regarding ankle and knee sprains in an attempt to explore the best way to summarize return and inform RTP decisions. The probability estimates will provide the predictive probability for how long until an athlete with varying demographics and injury characteristics will RTP. The use of the High School RIO™ ISS database for the purpose of evaluating RTP probabilities is approved under the University of Kentucky Institutional Review Board Exemption Certificate for Protocol #12-0409-X2H.
2 Return to Play after Sports Injury: A Review of the Literature

2.1 A decision-based return to play model

A predominant issue among sports medicine researchers related to return to play is the subjective nature in which RTP decisions are made. There is a high degree of variability among clinicians regarding factors considered for RTP [15], resulting in a call to develop an objective method to determine RTP. One validated decision-based model proposes a three-step process that requires the evaluation of health status (step 1), participation risk (step 2), and consideration of decision modifiers (step 3) [16, 17]. Health status of the athletes is based on medical factors such as patient demographics, symptoms, and clinical injury evaluation. These measures tend to be more objective relative to those in subsequent steps of the process. The evaluation of participation risk includes consideration of sport risk modifiers and provides more information about the type of sport, position played, competitive level, and the potential for utilization of protective equipment. However, general historical information about the injury is not considered in this step of the process. The consideration of decision modifiers allows for the influence of outside factors surrounding the athlete. Internal and external pressure [18], timing of the season, financial considerations, and fear of litigation are potential decision modifiers; the balance between risk and benefit of participation should always be considered when evaluating these factors.

Evaluation of health status is likely the first step taken by all athletic trainers and clinicians when assessing an injured athlete; however, this decision model provides specific factors to consider. Patient demographics and medical history, particularly regarding previous sports-related injuries to determine first or subsequent injury, provides
a context in which injury signs and symptoms can be evaluated [16]. In addition to pain, muscular strength and range of motion are the dominant health status factors assessed through functional tests to determine return to play potential; it is suggested that both be at or near preinjury levels before return is allowed, in a range of 70-100% [16, 19, 20]. The injury site should also be functionally stable with no tenderness, swelling, or effusion [16]. Girth should also be evaluated, although no criteria has been suggested for a RTP decision [16]. Appropriate laboratory tests should be conducted and reviewed to objectively evaluate tissue healing or identify physiological abnormalities if present [16]. Often overlooked, psychological state, in particular, readiness and confidence, is an important factor to consider when evaluating RTP [16]. Motivation during recovery has been shown to increase satisfaction with recovery outcomes [21]; apprehension and anxiety have been linked with higher rate of subsequent injury and shown to decrease performance [22].

The evaluation of participation risk relies on specific sports-related information. Type of sport, position played, and competition level should be considered [16]. In addition, the ability to protect the injury must also be evaluated [16]. This is not only related to protective equipment required for participation in the specific sport, but the ability to provide isolated protection to the injury itself. Taping, bracing, and splinting may be accommodated; however, athletes must adhere to the rules of their sport. This step is the most subjective of the three-step decision model as a standardized method to evaluate participation risk based on these sport-specific risk factors neither exists nor has been proposed. Further, participation risk should also be evaluated based on injury-specific risk factors. The risk of subsequent injury when a player is returned too early is certainly
important, but more essential, evidence regarding appropriate return timelines based on historical empirical evidence should be considered.

While the evaluation of participation risk is specific to the sport and injury, the final step of considering decision modifiers is the most specific to the individual athlete. Decision modification factors are those that are not related to health, sport, or injury, but can heavily influence a RTP decision. Both pressure from the athlete and external pressure from coaches, teammates, family, fans, and media can encourage an athlete to return to play too early. Time of season is suggested as a decision modifier [16], but it would have a better fit within the context of participation risk. For example, it could be argued that participation in an exhibition game would not pose as much of a risk to the athlete as participation in a play-off game. Other suggested decision modifiers are related to professional rather than recreational athletes. Conflict of interest, most common to paid clinicians, and fear of litigation for damages resulting from RTP too early [16] should be considered, but will likely not be factors for recreational athletes.

Validation of the proposed three-step RTP decision model indicated that in general sport participation restrictions increased as injury severity increased [17]. However, considerable heterogeneity in recommended restrictions was reported from the validation study, likely from varying interpretation of participation risk. This is further evidence indicating the need for objective methods to evaluate the risk of participation. This decision-based model is an important first step toward providing an ordered process in which clinicians can evaluate evidence leading to a RTP decision, and helps reduce the influence of clinical experiences. While the model does propose a more objective process, the evaluation of participation risk remains subjective in nature and lacks
empirical historical epidemiological data specific to the injury. If clinicians were provided probability estimates summarizing timeframes in which athletes with specific injuries return to play, they would have stronger evidence to make a determination of participation risk.

2.2 Return to play in sports medicine literature

The ability to return to play has long been a pivotal question posed by athletes and coaches after injury. Clinicians have relied on their personal experience to predict when an athlete might return, and RTP decisions vary tremendously between clinicians [15]. Sports medicine researchers have recently called for a consensus on RTP guidelines and criteria [23, 24], and researchers are beginning to answer that call.

2.2.1 Literature search methodology

Sports medicine literature published through October 2016 was searched using Medline through PubMed and SPORTDiscus and CINAHL through EBSCOhost. Search terms consisted of “return to play” in the title for all search engines and databases; for the CINAHL search, “sports medicine” was selected as the special interest. PubMed returned 306 results, SPORTDiscus returned 272, and CINALH returned 81. After removing duplicate publications, the remaining articles were manually evaluated for English language and relevance to return to play in athletes after injury. A total of 360 publications were identified for review; publications were categorized as books, reviews (book, clinical, comprehensive, narrative, literature, and systematic reviews and meta-analyses), editorials (editorial articles, comments, and conference proceedings), and original research. Publication identification is summarized in Figure 2.1.
2.2.2 Literature search results

To analyze the progression of sports medicine literature related to return to play, the trend in number of publications per year by publication type was evaluated. The earliest publication identified related to return to play was original research published in 1981 [25], detailing musculoskeletal profiling in terms of rehabilitation. There was only one other publication from the 1980’s, conference proceedings from 1984 discussing shoulder rehabilitation after rotator cuff tendonitis [26]. Return to play literature was published every year beginning in 1991; only publications identified between 1991 and 2016 were included when assessing trends over time. The identified publications address several specific types of injuries, as well as several different sports; these specifications are summarized in Table 2.1.

The number of articles within each publication type for each year from 1991 to 2016 are illustrated in Figure 2.2. Books, editorials, and reviews represent such a relatively small number of the total publications identified (collectively only 43 of 360 publications, or 12%) that an emerging trend is difficult to identify. However, it is clear that in general original research publications have seen an increase, particularly since 2010.

The original research publications were further evaluated to determine if they provided an outline for return to play criteria or guidance. Figure 2.3 illustrates the trend over time in number of publications providing RTP guidance, either to general athlete injuries, specific injuries, or specific sports. For ease of illustration, publications from 2015 and 2016 have been included with publications from 2010 through 2014 in this figure; 40 articles were published between 2010 and 2014 and 17 articles were published in 2015
and 2016. The number of publications outlining RTP criteria or guidance has been increasing since 1991, with a particularly large increase since 2010.

Evidence suggests that sports medicine researchers are now beginning to examine return to play after injury more thoroughly. Not only has the number of publications related to RTP increased, particularly since 2010, but the number of publications outlining RTP criteria or guidance on RTP has increased as well. Nearly all of the articles that provide estimates of return timelines present common epidemiological rates and proportions; the epidemiological incidence proportion is most commonly used and characterized as the probability of returning to play during the specified time interval \[27\]. In the context of RTP, the estimated epidemiological IP for RTP for a specific injury for a season is calculated as:

\[
IP = \frac{\text{number of athletes with injury that RTP during specified time interval}}{\text{number of injured athletes during specified time interval}}.
\]

The IP counts the number of injured athletes returning as opposed to the number of injuries. If an athlete suffers 2 lateral ankle sprains during the season and returns to play after both injuries, they would increase the numerator by 1, not 2. This provides a measure of the average “risk” of return to play for an athlete; however, as it does not distinguish athletes with single injuries from those with multiple or subsequent injuries in the calculation or interpretation, this can be a misleading illustration of the “risk” of return to play for an athlete.

This measure of incidence of RTP is accurate under the assumption that follow up is performed on all injured athletes and information regarding RTP is available. Lack of
follow-up outside of the season can lead to a violation of this assumption. If RTP status of the athlete is not known, athletes cannot be factored into the numerator as we cannot say whether the athlete returned to play. However, as the athletes are “at risk” for RTP, that is, they have suffered an injury, they would be included in the denominator. Not accounting for these cases would bias downward the estimates of T-RTP. Given the nature of the T-RTP outcome, it is likely that some athletes will not have a RTP event; cases in which there is no data available indicating whether an athlete returned to play can be considered censored, a common phenomenon that occurs when the time to event for an individual is only partially known [28-31]. In the context of return to play, a censored observation could occur in one of six ways: 1. the athlete was determined to be medically disqualified for the season, 2. the athlete was determined to be medically disqualified for their career, 3. the athlete chose not to continue but was not medically disqualified, 4. the athlete was released from the team but was not medically disqualified, 5. the athlete did not return for unspecified reasons, or 6. the season ended before the athlete could return to play. In all cases, the only known information about RTP is that it occurred sometime after the date of injury, indicating these observations are right-censored. These cases can be accounted for by applying time to event analysis methods. Five articles were identified that present RTP timelines using this analysis methodology [32-36].

2.3 Utilization of time to event analysis methods in current sports medicine literature

Use of time to event analysis methods to evaluate time to return to play after injury in sports medicine literature is relatively new and few studies have been published to date. To provide a significant contribution to the sports medicine literature, it is necessary to
provide a complete and accurate description of the current state of the use of this method
in the existing literature.

2.3.1 Literature search methodology

Sports medicine literature published in the English language through October 2016 was
searched using Medline through PubMed and SPORTDiscus and CINAHL through
EBSCOhost. Search terms consisted of “return to play” in the title and “time to return” in
the abstract for all search engines and databases; for the CINAHL search, “sports
medicine” was selected as the special interest. PubMed returned 23 results,
SPORTDiscus returned 26, and CINALH returned 7. These results were individually
evaluated to determine whether the methodology incorporated time to event analysis
applied to return to play; five articles were identified from this search procedure [32-36].

2.3.2 Literature search results

Five articles using time to event analysis methodology were published between 2012 and
2015. The two most recent publications used time to return to play as the outcome, but
focused on comparing two different treatment methods. One article compared time to
return to play in a randomized, three-arm parallel-group trial for treatment of acute
hamstring injury using platelet-rich plasma injection [35]. The other article compared
time to return to play in elite professional soccer players between those undergoing
lateral versus medial meniscectomy surgeries [36]. The three remaining publications use
time to event analysis to summarize return times after specific injury. The earliest article
develops predictive linear and Cox regression models for time to return to play after high
ankle sprain in a convenience sample of 20 college football players [34]. From the Cox
model, hazard ratios are presented comparing RTP between linemen and other positons,
as well as the association between injury severity measured as height of tenderness and RTP. This article does not provide any RTP timelines or probabilities, nor does it detail components of the Cox model, in particular, censored cases of return. While not explicitly stated, it is inferred that T-RTP measured as a continuous variable is the outcome variable for the Cox model.

The final two articles, published by the same author, present return to play probabilities and timelines in high school athletes; one paper focuses on concussions [33] and the other compares new versus recurrent ankle sprains while providing RTP probabilities for both [32]. In both cases, T-RTP was defined as time lost from participation and measured as an ordinal categorical variable: same-day return, 1-2 day (next day) return, 3-6 day return, 7-9 day return, 10-21 day return, and >21 day return. However, an additional category was added called “no return [censored data]” that transforms T-RTP to a nominal categorical variable. There are two potential issues with this categorization of the T-RTP variable. First, coding RTP as a categorical variable essentially creates an arbitrary ordered outcome variable of time; calculating Kaplan-Meier probabilities for the ordinal categories does not account for how the intervals are actually defined. It is best to measure time to return to play as accurately as possible; however, both of these papers use data collected on a standardized injury report form where time lost from participation was collected categorically. In this situation, it is more appropriate to estimate RTP probabilities using the Life-Table Method as this is better suited to calculate estimates for time in intervals [31, 37]. Second, the upper tail of the survival distribution is poorly estimated when a sizable number of the cases are censored [31]. Including all censored cases in the largest event time provides unstable estimates of survival probabilities in the
upper tail of the distribution and implies censoring times are greater than the largest event time, which will bias the mean survival time downward.

An alternate method of incorporating the censored cases in the analysis is to classify those cases into the most appropriate return category and mark them as censored. If the date of injury is known, the time between injury and last contact with the athlete (e.g., the end of season date) could be calculated and categorized appropriately. If time is measured as a continuous variable, the difference in last contact time and injury time would be used. Further, it is suggested that standardized injury report forms used by athletic trainers allow time lost from participation to be collected as accurately as possible. Fields could be included that capture injury date and time as well as date and time an athlete is cleared to return to participation.

In both articles, Kaplan-Meier estimates are calculated and subtracted from 1 and reported as RTP probabilities for each time interval. The KM estimates, \( \hat{S}(t_j) = P(\text{TRTP} \geq t_j) \), are the probability that RTP took place in the time interval \( t_j \) or later. The reported estimates, \( 1 - \hat{S}(t_j) = \hat{F}(t_j) = P(\text{TRTP} < t_j) \), are the probability that RTP took place prior to the time interval \( t_j \). If T-RTP is a discrete random variable that takes values \( t_1 < \cdots < t_k \), the probability density function, \( f(t_j) = P(\text{TRTP} = t_j) \), is the probability that RTP takes place in the time interval \( t_j \) and the hazard at time \( t_j \), \( \lambda(t_j) = P(\text{TRTP} = t_j|\text{TRTP} \geq t_j) \), is the conditional probability of RTP in time interval \( t_j \), given that an injured athlete has not returned to play at that point. It is the pdf that provides the probability of RTP in each interval, but the hazard provides the probability of RTP in each interval accounting for censored cases. It would be more accurate to report the
estimated hazard as the probability of RTP in each time interval; these estimates are easily calculated by applying the Life-Table Method.

Aside from the issues regarding measurement of T-RTP and calculation of RTP probabilities, these two publications are a significant step toward providing evidence-based, objective prognostic indicators of when an athlete is likely to return to participation. Further, both illustrate the necessity of providing guidance to sports medicine researchers with reference to applying and interpreting time to event analysis methods to return to play.

2.4 Epidemiology of ankle and knee sprains and return to play

Risk factors for ankle and knee injuries in athletes have been extensively documented in sports medicine literature [2, 38-56]. Less documented, however, are factors that influence RTP timelines after lateral ankle and knee ligament sprains. An exploration of the epidemiology of these sports injuries and potential factors affecting RTP is necessary before attempting to evaluate multivariate T-RTP relationships.

2.4.1 Lateral ankle sprain

Ankle sprains are the most common lower extremity (LE) orthopaedic injury, with approximately 23,000 ankle sprains occurring daily in the U.S. [2, 57]. An estimated 1.6 million physician office visits and over 8,000 hospitalizations per year are attributable to ankle or foot sprains [58], and associated healthcare costs have been estimated at 4.2 billion dollars per year in the U.S. alone [59, 60]. Ankle ligament sprains are the most common injury across field and court sports, accounting for anywhere from 15% to 75% of all reported injuries [1, 8]. Acute lateral ankle sprains are common among young
athletes under age 18, occurring with the foot is plantar flexed and inverted [7]. Incidence of ankle sprain has been shown to be higher in adolescents than adults [51]. Men’s and women’s basketball maintain the highest ankle sprain rates [51, 61], and along with women’s outdoor track, women’s field hockey, and soccer, maintain the highest injury recurrence rates as well [8, 61]. Findings have suggested that incidence of ankle sprain is higher in females than males [61]; however, other studies have shown sex does not appear to be a risk factor [40]. Similarly, researchers have failed to reach a consensus on whether height, weight, limb dominance, muscle strength, muscle reaction time, and postural sway are potential risk factors for ankle sprain [40].

Ankle sprains can lead to residual impairments such as re-sprain, perceived instability, functional instability, mechanical instability (joint laxity), pain, swelling, a feeling of weakness, and subsequently reduced level of physical activity [62]. Suffering one or more of these residual impairments is known as chronic ankle instability (CAI). Approximately 30% of those who suffer a first-time ankle sprain develop CAI, although this has been reported as high as 70% [8, 14, 63]. Functional testing is necessary throughout the rehabilitation process to objectively gauge the athlete’s progress in regaining balance, proprioception, strength, range of motion, and agility [8]. In addition to the high incidence rate, ankle sprains also have a high rate of recurrence, particularly when athletes return to play too early. Among NCAA athletes, 1 in 8 ankle sprains was identified as recurrent [61]. Basketball athletes are 5 times more likely to experience subsequent ankle sprains; the injury recurrence rate has been reported as high as 73% [8]. It has been reported that across all sports, once an ankle sprain occurs, up to 80% will suffer subsequent sprains [8]. Ankle sprain rates have been shown to be higher in
competition versus practice [1, 8, 61], and preseason practice injury rates have been reported higher than those of in-season and post-season [1].

In NCAA athletes, nearly half of ankle sprains were non-time loss (NTL) injuries in which RTP occurs within 24 hours after injury; nearly 5% of ankle sprains required more than three weeks before RTP, including those who did not RTP at all [61]. Patients treated for acute lateral ankle sprains have shown decrease in pain and improved motion and function within 2 weeks of injury; however, it has been reported that 5-25% of injured athletes were still experiencing pain or occasional instability at 1 year [64]. Further, more than one-third have reported reinjury at 3 years [64].

A combination of subjective and objective indicators is necessary to accurately determine when athletes can safely RTP following ankle sprain. While foot and ankle scoring systems do exist, none have been validated for RTP decisions [8]. There is a lack of evidence-based guidelines for RTP decisions after ankle injury; this void creates a challenge in determining the acceptable time in which an athlete may safely RTP [8]. While researchers have not yet come to a consensus on potential risk factors for lateral ankle sprain [42], they do agree the primary predisposing factors for experiencing an ankle sprain is a history of previous sprains [40, 42] and premature RTP [8]. When determining factors that could potentially affect T-RTP, risk factors for injury should be considered. Agreement on history of previous sprain, time in season, and type of exposure (competition versus practice) indicates these factors should be considered. The proposed 3-step model for RTP determination recommends the consideration of health status and participation risk including pain, muscle strength, range of motion, psychological status, type of sport, injury protection, and time of season [16]. Within the
HS RIO™ database, basic demographics (gender, age, year in school), competition characteristics (time in season, competition versus practice, competition site, and competition time), and injury descriptions are recorded. All of these factors should be considered for analysis. Restricting analysis to only field and court sports will control for some of the observed difference in injury risk between sports. While the lack of previous ankle sprain and individual health status information is a limitation and should be considered in future studies.

2.4.2 Knee ligament sprain

Knee sprains are the second most common LE sports-related injury [2, 3, 65]. Medial collateral ligament (MCL) sprains are the most common ligament injuries at the knee and occur frequently in football [66], ice hockey [67], soccer [68], and skiing [69], all sports in which body movements causing high valgus stress at the knee are common [70, 71]. Incidence of anterior cruciate ligament (ACL) injury has recently risen significantly in college and adolescent athletes [1, 7]. This rise could potentially be due to improvements in identification of injury through diagnostic testing [1], or an increase in participation in sports where the mechanisms of ACL injury are common. Basketball, football, and soccer participation has increased among adolescent athletes; these sports often require deceleration or change of direction forces [7]. ACL injuries typically occur in non-contact conditions [7], although the opposite has been reported for male athletes [38]. Girls are more susceptible to ACL injury than boys, although the underlying cause of this increased vulnerability is still unclear [7, 40]. Theories posit that girls participate in less strength and conditioning, have a smaller ligament with a smaller intercondylar notch at the femur, have different mechanics during play, and have a different anatomical
alignment [7]. Increased body mass index (BMI) has been observed to be a risk factor for ACL injuries, particularly among females; however, several studies have found no impact on injury risk due to BMI [72]. Generalized joint laxity and small and weak ACL have been identified as risk factors for ACL injury [72]. ACL injury rates for male athletes are highest for football, both in competition and practice [1]; similar results have been found for other sports as well [73]. For female athletes, the highest ACL injury rates are reported in lacrosse [38].

Previous ACL injury has not been identified as a risk factor for subsequent injury in male athletes; however, for female athletes, subsequent ACL injury is a risk factor for future ACL surgery, and ACL reconstruction on the non-dominant knee is a risk factor for future ACL injury [73]. Severity of knee sprain is a subjective measure by both the practitioner and athlete, and too unreliable for inclusion in analysis. As knee injuries typically involve multiple ligaments, the number of injured ligaments could be used as a proxy for injury severity; however, number of ligaments injured has no effect on RTP [74]. Similar to ankle sprains, knee sprains lead to time lost from activity, functional instability, chronic instability, and joint degeneration over time even though surgery is not typical. For mild MCL sprains, reported diminished functional capacity lasts for several weeks with RTP ranging from 4 to 19 days post-injury [75]; reported diminished functional capacity for moderate MCL sprains ranges from 3 to 8 weeks [76]. ACL RTP guidelines suggest it could take between 4 and 8 weeks for full range of motion to return and swelling to subside [77]. Mild to moderate LCL sprains can heal within 2 to 4 weeks [78]. As evidenced by these imprecise timeframes, healing and RTP timelines are difficult to predict based on tissue damage alone.
There is a lack of evidence concerning risk factors for single-ligament LCL and PCL injuries in the existing literature; however, differences in risk factors and injury incidence rates between ACL and MCL indicate the necessity to stratify analyses by knee ligament injured. Similar to lateral ankle sprains, there is a lack of evidence-based guidelines for RTP decisions after knee ligament injury, providing a challenge for determining the acceptable time in which an athlete may safely RTP. Researchers have agreed upon two risk factors for knee ligament injury: being a female athlete and previous knee injury [79]. Agreement on history of previous sprain and gender indicates these factors should be considered in analyses of T-RTP. All demographic, competition, and injury factors available in the HS RIO™ database should be considered and the lack of previous injury and individual health status information noted as a limitation of analysis.
<table>
<thead>
<tr>
<th>Type of Injury</th>
<th>n (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Head/concussion/face</td>
<td>98 (27.2)</td>
</tr>
<tr>
<td>Neck/cervical spine</td>
<td>37 (10.3)</td>
</tr>
<tr>
<td>Shoulder</td>
<td>18 (5.0)</td>
</tr>
<tr>
<td>Arm/elbow/wrist/hand</td>
<td>15 (4.2)</td>
</tr>
<tr>
<td>Hip/trunk</td>
<td>6 (1.7)</td>
</tr>
<tr>
<td>Hamstring</td>
<td>23 (6.4)</td>
</tr>
<tr>
<td>Knee</td>
<td>35 (9.7)</td>
</tr>
<tr>
<td>Leg/quadriceps/Achilles/ankle/foot</td>
<td>20 (5.6)</td>
</tr>
<tr>
<td>Musculoskeletal</td>
<td>7 (1.9)</td>
</tr>
<tr>
<td>Muscle/soft tissue</td>
<td>13 (3.6)</td>
</tr>
<tr>
<td>Cardiac event (acute)</td>
<td>10 (2.8)</td>
</tr>
<tr>
<td>Abdomen (internal)</td>
<td>6 (1.7)</td>
</tr>
<tr>
<td>Circulatory/respiratory/thyroid</td>
<td>4 (1.1)</td>
</tr>
<tr>
<td>Heat stroke</td>
<td>4 (1.1)</td>
</tr>
<tr>
<td>Infectious disease</td>
<td>4 (1.1)</td>
</tr>
<tr>
<td>Mental health</td>
<td>3 (0.8)</td>
</tr>
<tr>
<td>Pregnancy/female athlete triad</td>
<td>2 (0.5)</td>
</tr>
<tr>
<td>Not specified</td>
<td>55 (15.3)</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Sport</th>
<th>n (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Baseball</td>
<td>9 (2.5)</td>
</tr>
<tr>
<td>Basketball</td>
<td>5 (1.4)</td>
</tr>
<tr>
<td>Football</td>
<td>39 (10.8)</td>
</tr>
<tr>
<td>Golf</td>
<td>1 (0.3)</td>
</tr>
<tr>
<td>Hockey</td>
<td>9 (2.5)</td>
</tr>
<tr>
<td>Karate</td>
<td>1 (0.3)</td>
</tr>
<tr>
<td>Rugby</td>
<td>7 (1.9)</td>
</tr>
<tr>
<td>Soccer</td>
<td>9 (2.5)</td>
</tr>
<tr>
<td>Swimming</td>
<td>1 (0.3)</td>
</tr>
<tr>
<td>Tennis</td>
<td>1 (0.3)</td>
</tr>
<tr>
<td>Track and Field</td>
<td>1 (0.3)</td>
</tr>
<tr>
<td>Not specified</td>
<td>277 (76.9)</td>
</tr>
</tbody>
</table>
Figure 2.1. Inclusion criteria and flow of retrieved articles

- Potentially relevant literature (n=659)
  - duplicates (n=231)

- Literature retrieved for more detailed evaluation (n=428)
  - non-English and irrelevant (n=68)

- Literature included (n=360)
  - original research (n=317)
  - editorials (n=11)
  - reviews (n=21)
  - books (n=11)

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Figure 2.2 Total number of publications by publication type

- Book
- Editorials
- Original research
- Review

Year:
- 1991
- 1993
- 1995
- 1997
- 1999
- 2001
- 2003
- 2005
- 2007
- 2009
- 2011
- 2013
- 2015

Number of Publications:
- 0
- 10
- 20
- 30
- 40
- 50
- 60
Figure 2.3 Number of original research publications providing RTP criteria or guidance
3 An Innovative Methodological Approach in Sports Medicine Research: Applying Time to Event Analysis to Return to Play

3.1 Introduction

Return to play (RTP) has historically been determined using subjective reasoning; there is a need for more objective methods to assist in the determination of RTP. Typical research studies involving RTP report rates or proportions, but these measures can be inaccurate in instances where a lack of follow-up data excludes some athletes from analysis. Time to event analysis is a commonly used statistical analysis method that can provide more accurate estimates for time to return to play (T-RTP) by accounting for all injured athletes regardless of lack of follow up concerning RTP. These analysis techniques have not been heavily applied or related to RTP, and will provide athletic trainers, coaches, and team physicians a more accurate way to estimate return time. For prognosis, it is better to know the likelihood of when an individual will experience the outcome of interest as opposed to summary results from cumulative risks and rates. In addition, it is likely that some athletes will not experience a return to play event, indicating the presence of censored cases. Time to event analysis can be applied to address both issues by generating evidence-based, objective estimates of when an athlete is likely to return following a given injury while accounting for censored cases. The accuracy of predicting time to return to play can be improved using these techniques, resulting in better patient care through education, improved coach-medical staff relations, and more efficient use of an athletic trainer’s clinical time. However, it is important that sports medicine researchers have a fundamental understanding of time to event methodology to
appropriately conduct analysis of return to play as well as interpret results and translate results to a clinical setting.

3.2 Background

Few studies have been conducted using time to event analysis methods to provide estimates of RTP probabilities. A survey of existing sports medicine literature identified only two publications providing RTP estimates; one paper focused on concussions [33] and the other compared new versus recurrent ankle sprains [32]. It is common for RTP data to be measured as an ordinal categorical variable and is often collected on standardized injury report forms as same day return, 1-2 (next day) return, 3-6 day return, 7-9 day return, 10-21 day return, and >21 day return. Due to the discrete nature of T-RTP and the unequal time intervals, RTP probabilities can be best estimated by applying the nonparametric Life-Table Method. In both publications, Kaplan-Meier (KM) estimates are subtracted from 1, resulting in the failure probability, and presented as RTP probabilities. KM survival probability estimates are interpreted as the probability that RTP took place in a specified time interval or later; failure probability estimates are the probability that RTP took place prior to a specified time interval. In the discrete setting, the probability density function provides an estimate of the probability that RTP takes place in a specific time interval, and the hazard function is the conditional probability that RTP takes place in a specified time interval given that an injured athlete has not returned to play before that time interval. Therefore it is not the failure probability that provides an estimate of when an athlete will return to play; it is the hazard probability that provides this estimate while accounting for censored cases. Functions necessary for the application of time to event analysis methods are detailed in Table 3.1 [28, 30, 31, 37, 80]. Note that
although time to return to play can be measured as a continuous or discrete variable, the
discrete case is provided here as it is more common for T-RTP.

### 3.3 Time to return to play analysis method

**Time to Return to Play Outcome.** The biggest distinction between time to event
analysis and other methods is the unique waiting time outcome variable, in this case time
to return to play. T-RTP is defined as the time between injury and when the athlete is able
to return to play. The waiting time outcome variable contains two parts: (1) time to return
to play and (2) an indicator for the occurrence of the event. T-RTP can be measured as a
quantitative or ordinal categorical variable. When measured as a quantitative variable,
return time is calculated as the difference between injury date and return date. We cannot
assume normality of waiting time outcomes [81]; the distribution is likely to be right-
skewed as most athletes are likely to RTP relatively quickly but some injured athletes
may take longer to return. Ideally the outcome variable should be measured as precisely
as possible; however, it is more common for T-RTP to be measured in categorical
intervals. The indicator for the occurrence of the event is a dichotomous variable that
takes a value of 1 if the athlete returns to play and a value of 0 if the athlete does not
return to play. It is important that the study period be sufficiently long to allow athletes
an opportunity to experience the event of RTP. At the end of the study period, all athletes
who have not returned to play are considered censored cases. Due to the nature of
surveillance for sports injuries, the follow-up period typically ends at the end of the
season. This lack of follow-up information for RTP can result in a high number of
censored cases.
Censored Cases of Return to Play. Athletes who do not return to play after injury are considered censored cases of RTP. There are three different types of censoring: right censoring, left censoring, and interval censoring [31]. An observation is right censored if it is only known that the time to event is greater than some value. This is the most common form of censoring as a study may end before the event occurs. In the context of RTP, right censoring could occur if an athlete is injured during the season but the season ends before they are cleared to RTP. An observation is left censored if it is only known that the time to event is less than some value. This could occur if, for example, an athlete had an ankle sprain at the start of the season, that is, the injury occurred before the observation period began. An observation is interval censored if it is only known that time to event is between two values. For example, consider evaluating RTP in boys’ soccer, which has a fall and a spring season. Suppose an athlete is injured during the fall season and does not return before season end, but is cleared to return before the spring season starts. The exact time of return to play is not known, only that it occurred sometime between the last day of the fall season and the first day of the spring season.

For return to play, a censored observation could occur in one of six ways: 1) the athlete was determined to be medically disqualified for the season; 2) the athlete was determined to be medically disqualified for their career; 3) the athlete chose not to continue but was not medically disqualified; 4) the athlete was released from the team but was not medically disqualified; 5) the athlete did not return for unspecified reasons; or 6) the season ended before the athlete could return to play. In all cases, no data will be available for RTP as the return time for the athlete is missing. Since all that is known about RTP is that it occurred sometime after injury, these observations are right-censored. Further, all
scenarios indicate non-informative cases of censoring as none are directly related to the study itself, that is, a censored case at a specified time point is representative of all other cases that have not experienced the event up to that time point [81]. T-RTP must be computed for censored cases of RTP as these athletes have not experienced the event and will have missing values of the waiting time outcome variable; censored cases will have a value of 0 for the event indicator variable.

**Life-Table Method for estimating hazard.** The LIFETEST procedure with the METHOD=LIFE and INTERVALS options specified in SAS will generate life-tables for RTP. Comparisons between strata can be analyzed using log-rank tests invoked by the STRATA statement. Adjustments for multiple comparisons can be applied by using the ADJUST option in the STRATA statement. Life-tables can be generated in SPSS through the Life Tables dialogue box by clicking Analyze -> Survival -> Life Tables. Comparisons between “By Factors” (strata) can be analyzed by specifying “Compare” options.

**Discrete logistic regression model for estimating hazard odds ratios.** The life-table method is useful for estimating RTP probabilities, conducting exploratory data analysis, and evaluating differences in survival curves across strata. If it is of interest to investigate multiple-variable relationships between T-RTP and injury, competition, and athlete demographic characteristics, Cox proportional-hazards regression models should be used. However, due to the discrete nature of T-RTP, the discrete logistic regression model for discrete time data should be used to estimate discrete-time hazard odds ratios [28, 31, 37].
The discrete logistic regression model, a proportional odds model, can be used to estimate the discrete-time hazard, \( P_{it} = P(RTP_i = t | RTP_i \geq t) \), the conditional probability that an individual \( i \) will RTP at time \( t \), given that individual has not already returned to play [31, 37]. The discrete logistic regression model for discrete time data uses the logit, or hazard odds, of \( P_{it} \) and takes the following form:

\[
\log\left[ \frac{P_{it}}{1 - P_{it}} \right] = \alpha_t + \beta_1 x_{i1} + \cdots + \beta_k x_{ik}.
\]

The parameter estimates provide estimates of the log hazard odds of RTP [82]. For dichotomous independent variables, \( \exp\{\beta\} \) is the hazard odds ratio; for continuous independent variables, \( 100 \times [\exp\{\beta\} - 1] \) gives the estimated percent change in the hazard odds for each one-unit increase in the covariate. This model can be estimated using the partial likelihood method, where the \( \alpha_t \)'s are treated as nuisance parameters and only \( \beta \)'s are estimated [31, 37]. The discrete logistic regression model can be estimated using the PHREG procedure in SAS and specifying the TIES=DISCRETE option in the MODEL statement.

Computing time using the partial likelihood can be large for large datasets with a high number of tied event times. To reduce computing time, the model can be estimated using the maximum likelihood method, which uses the full likelihood to explicitly estimate both the \( \alpha_t \)'s and the \( \beta \)'s [31]. This allows for direct hypothesis testing regarding changes in the hazard over time that is not possible using the partial likelihood. For this method, a new dataset is generated based on the original data containing pseudo-observations, one for each time category of follow-up for each individual, with a variable indicating whether the event has been experienced at that time point. A logistic regression model is
then fit by the LOGISTIC procedure in SAS using the pseudo-event indicator as the response variable and pseudo-event time as an independent variable along with the other covariates.

3.4 T-RTP example

An observational study is used to illustrate the application of time to event analysis methods to return to play. Analysis methods will be presented for SAS (version 9.4, SAS Institute, Inc., Cary NC, USA), and SPSS (IBM SPSS Statistics for Windows, Version 23, IBM Corp., Armonk, NY, USA).

3.4.1 Design and sample

A secondary data analysis of athlete injury data from the High School RIO™ Injury Surveillance System (ISS) database was conducted. All US high schools with a National Athletic Trainers’ Association (NATA) certified athletic trainer (AT) were eligible for enrollment in this ISS; AT’s who enrolled their school completed the online “Exposure Report Form” for reportable injuries each week. New lateral ankle sprains experienced by high school athletes from approximately 100 high schools in the US during regularly scheduled participation in school-sanctioned sports for seven academic years (2005-2006 through 2011-2012) were used in this illustration. Field and court sports (football, boys/girls soccer, volleyball, wrestling, basketball, baseball, and softball) were considered as these athletes are more likely to suffer lateral ankle sprain. The use of the HS RIO™ ISS database for the purpose of evaluating RTP probabilities was approved by the University Institutional Review Board.
Within the HS RIO™ ISS, a sprain was defined as injury to the ligamentous or capsular tissue [2, 83]. All ankle sprains that required the athlete to be removed from participation and diagnosed as an injury by the treating health care professional were reported, regardless of time lost from participation. No unique personal identifying information was contained in the dataset; the de-identification process of the data prohibits linking multiple injuries on the same athlete allowing only new injuries to be considered for analysis. A new injury was defined as an ankle sprain with an acute, traumatic onset of symptoms with no prior history of that injury. Ligament damage is likely the strongest indicator of severity; LAS were graded on the number of lateral ligaments that were damaged [84]. The ligaments under consideration include the anterior talofibular (ATF), calcaneofibular (CF), and posterior talofibular (PTF) [85, 86]. An injury with one-ligament damaged was classified as a first degree sprain and considered mild; two ligaments was a second degree sprain and considered moderate; three ligaments was a third degree sprain and considered severe. For lateral ankle sprain, ligament healing times may not be a strong indicator for when an athlete will RTP. Although many RTP decisions are centered on symptom resolution clearing athletes to RTP prior to complete tissue healing, the extent of tissue damage may still contribute to RTP timelines. RTP probabilities are presented stratified by severity for new LAS.

3.4.2 Measures

**Time to Return to Play Outcome.** In the HS RIO™ ISS database, the number of days the athlete was withheld from participation was collected in intervals of 1-2 days, 3-6 days, 7-9 days, 10-21 days, and more than 21 days [83]. If the athlete had a reported
return time, the event indicator variable was given a value of 1; otherwise the event indicator was given a value of 0.

**Censored Cases of Return to Play.** Estimated season end dates for each sport were used as a proxy for last date of contact. The number of days between injury date and season end date was calculated, and each injury was then classified into the appropriate T-RTP category. All censored cases were assigned an event indicator of 0. Estimated season start and end dates for each sport are listed in Table 3.2.

**Athlete Demographics and Competition Characteristics.** Demographic characteristics were documented for each injured athlete including year in school (freshman, sophomore, junior, senior), age in years, gender, height in inches, weight in pounds. Competition characteristics at the time of injury were also documented including sport in which the athlete was participating, time in season (preseason, regular season, postseason), type of exposure (competition or practice), competition site (home, away, neutral site), and competition time (warm-ups, beginning, middle, end, overtime). Indicators for sport (football, soccer, volleyball, wrestling, basketball, baseball/softball) were included to account for differences in injury frequency among field and court sports that may influence RTP.

3.4.3 Data analysis

**Life-Table Method for estimating hazard.** T-RTP is a discrete random variable that can take on values $t_j, j = 1,2,3,4,5$, where $t_1 < t_2 < t_3 < t_4 < t_5$. For each injury recorded, the data set must include the categorical T-RTP interval and event indicator variables, as well as a measure of 1<sup>st</sup>, 2<sup>nd</sup>, and 3<sup>rd</sup> degree ankle sprain as a proxy for severity. The first
20 records of the dataset are printed in Figure 3.1. The LIFETEST procedure in SAS will generate a life-table for RTP:

```sas
PROC LIFETEST DATA=ankle METHOD=LIFE INTERVALS=(1,2,3,4,5);
   TIME trtp * event(0);
   STRATA ankle_sprain_severity;
RUN;
```

The METHOD=LIFE option must be specified to create a life-table. By default, SAS will attempt to create 10 intervals from the time variable; in the case of RTP data, we must specify interval length of 1 since our T-RTP variable takes on categorical values 1 through 5. The TIME statement contains the T-RTP variable and an indicator for censored cases of RTP. The STRATA statement is used to generate RTP probabilities by severity. The life-table for 1st degree ankle sprains is provided in Figure 3.2.

For RTP, we disregard the first row of the output as it has no practical interpretation. The next five rows of output provide life-table estimates for 1-2 day return, 3-6 day return, 7-9 day return, 10-21 day return, and more than 21 day return. The summary table reports 1,305 1st degree ankle sprains documented for academic years 2005-2006 through 2011-2012. The “Number Failed” column is the number of athletes who returned to play during that interval, $d_j$. The “Number Censored” column provides the number of censored cases for that interval, $w_j$. Not included in the output table is the number of athletes entering the interval without yet experiencing RTP, $Y'_j = Y'_{j-1} - d_{j-1} - w_{j-1}$. The “Effective Sample Size” column is the number of athletes “at risk” for return during that interval, $Y_j = Y'_j - \frac{w_j}{2}$. Since these censored cases are only at risk for half of the interval, they only count for half of the interval when calculating the effective sample size. The “Conditional Probability of Failure” is the conditional probability of an injured athlete experiencing
RTP in the interval given that they have not experienced RTP before that interval, that is,

\[ P(\text{TRTP} = t_j \mid \text{TRTP} \geq t_j) \], calculated by \( \frac{d_j}{y_j} \). This is equivalent to the discrete time hazard. An estimate of its standard error is given in the next column.

Survival and failure probabilities are calculated by the following formulas:

\[ \hat{S}(t_j) = P(\text{TRTP} \geq t_j) = \prod_{i=1}^{j} \left( 1 - \frac{d_{i-1}}{y_{i-1}} \right) \text{ where } \hat{S}(t_1) = 1 \]

\[ \hat{F}(t_j) = P(\text{TRTP} < t_j) = 1 - \hat{S}(t_j). \]

The standard error of the survival probability estimates is provided in the next column. The median residual lifetime is an estimate of the remaining T-RTP for an athlete that has not yet returned at the start of the interval. Both pdf and hazard estimates evaluated at the midpoint of the interval are provided, along with their standard errors. The intervals are notated in the output as \([\text{lower}, \text{upper})\). The estimates of the pdf and hazard for each interval are calculated using the following formulas:

\[ \hat{f}(t_j) = P(\text{TRTP} = t_j) = \frac{\hat{S}(a_{j-1}) - \hat{S}(a_j)}{\text{upper} - \text{lower}} \]

\[ \hat{h}(t_j) = P(\text{TRTP} = t_j \mid \text{TRTP} \geq t_j) = \frac{d_j}{(\text{upper} - \text{lower})(y_j - \frac{d_j}{2})}. \]

By definition, the discrete-time probability density function is the probability that an athlete will return to play at a specified time interval \( t_j \). However, the more accurate estimate of RTP is the hazard, defined as the conditional probability of RTP at a specified time interval \( t_j \), given that the athlete has not returned to play before \( t_j \), as it accounts for
censored cases of RTP. Notice that the estimate of the hazard function as defined by SAS is similar to the conditional probability of failure, but accounts for the length of the interval and evaluates at the midpoint by removing half of the “failures” from the denominator. The interval lengths we have designated are arbitrary, so accounting for interval length and evaluating at the midpoint does not offer any practical benefit.

Although labeled in the SAS output as the hazard, the column labeled conditional probability of failure provides the discrete time hazard probability of RTP for each time interval.

To more accurately reflect the intervals in which RTP is collected, the TRTP variable and the INTERVALS option in PROC LIFETEST could be modified. The TRTP variable would need to be defined as the left endpoint of the RTP intervals; instead of categories of intervals 1, 2, 3, 4, and 5 (time_to_return), they would be 1, 3, 7, 10, and 22 (time_to_return_c). The INTERVALS option and TIME statement would be coded as:

```plaintext
PROC LIFETEST DATA=ankle METHOD=LIFE INTERVALS=(3,7,10,22);
  TIME trtp_c * event(0);
  STRATA ankle_sprain_severity;
RUN;
```

This would only affect the values of the upper and lower endpoints, resulting in different estimates of the pdf and hazard evaluated at the midpoint of the interval. This does not change the conditional probability of failure estimates that are most applicable to athletes and those interested in RTP. The option to specify unequal discrete time intervals is not available in SPSS software.

Using the same set of data illustrated in Figure 3.1, RTP probability estimates can be calculated using SPSS. The Life Tables dialogue box can be accessed by clicking
Analyze -> Survival -> Life Tables… and making the selections in Figure 3.3. After “event” has been moved to the Status box, click the Define Event… box, select the Single value radio button, and enter “1” as this indicates the athlete has experienced the event. After “ankle_sprain_severity” has been moved to the By Factor box, click the Define Range… box and enter the minimum (1) and maximum (3) values for ankle sprain severity.

Comparisons between “By Factors” (strata) can be analyzed by clicking the “Options” box in the upper right corner of the Life Tables dialogue box and selecting “Overall” under “Compare Levels of First Factor.” Selecting “Pairwise” will provide multiple comparisons between strata.

The Life-Table output from SPSS is provided in Figure 3.4. All of the columns in this table align with those provided in SAS. In particular, the “Proportion Terminating” column in SPSS aligns with the “Conditional Probability of Failure” column in SAS and provides RTP probability estimates by severity for ankle sprains.

Common plots can be selected using the PLOTS option in the PROC LIFETEST statement in SAS or in the Options box in SPSS. Plots of survival, log survival, pdf, and hazard versus time are available in both software packages; however, there is no direct option for a plot of RTP probability estimates against time.

**Discrete logistic regression model for estimating hazard odds ratios.** The discrete logistic regression model is estimated using the PHREG procedure in SAS and specifying the TIES=DISCRETE option in the MODEL statement:
PROC PHREG DATA=ankle;
   CLASS ankle_sprain_severity;
   MODEL trtp * event(0) = ankle_sprain_severity / TIES=DISCRETE;
RUN;

To avoid computing time constraints in datasets with high numbers of tied event times, this model can be estimated with the LOGISTIC procedure using the pseudo-dataset created with the following SAS code:

DATA ankle_pseudo;
   SET ankle;
   
   DO time_int=1 TO trtp;
      IF time_int = trtp AND event = 1 THEN rtp = 1;
      ELSE rtp = 0;
   OUTPUT;
   END;
RUN;

A comparison of the original ankle dataset and the dataset consisting of pseudo-observations for the first six athletes is provided in Figure 3.5. A logistic regression model is fit using the pseudo-event indicator as the response variable and pseudo-event time as an independent variable along with the other covariates. This model provides estimates of the effect of time on the hazard odds of returning to play controlling for ankle sprain severity, athlete demographics, and competition characteristics. Competition site and competition time were not included in the model due to excessive missing values (86% missing for both variables). Because we are interested in the differences in RTP play probabilities by ankle sprain severity, we include the interaction between time and severity. The model is produced using the LOGISTIC procedure in SAS:

PROC LOGISTIC DATA=ankle_pseudo;
   CLASS time_int ankle_sprain_severity year_in_school gender
time_in_season type_of_exposure soccer volleyball wrestling
basketball baseball_softball / PARAM=GLM;
MODEL rtp(DESC) = time_int ankle_sprain_severity 
    time_int * ankle_sprain_severity 
    year_in_school gender height_inches weight_pounds 
    time_in_season type_of_exposure soccer volleyball 
    wrestling basketball baseball_softball / LACKFIT;
RUN;

All independent variables included in the CLASS statement will be treated as categorical. The PARAM=GLM option in the CLASS statement overrides the default effect coding in the design matrix and creates a set of indicator variables with the highest value of the categorical variable as the reference. The DESCENDING option in the MODEL statement will predict the probability of rtp=1 instead of rtp=0. The LACKFIT option in the model statement provides the Hosmer and Lemeshow goodness of fit test for the logistic regression model. The SAS output for this model is provided in Figure 3.6.

Notice in the Type 3 Analysis of Effects that the interaction between time and severity is significant \( (p=.001) \). This means that the hazard odds of RTP varies for different time intervals and levels of severity. The LSMEANS statement in PROC LOGISTIC can be used to calculate hazard odds ratios for the different combinations of time and severity:

\[ \text{LSMEANS } \text{time_int} * \text{ankle_sprain_severity} / \text{PDIFF EXP CL}; \]

The PDIFF option performs comparisons of LS Means, and the EXP and CL options provide the hazard odds ratios and confidence limits for the different combinations of time and severity.

3.4.4 Results

There were 2,086 documented new lateral ankle sprains among field and court sport athletes with a known injury date. Of those, 1,305 (62.6%) were 1st degree sprains, 645 (30.9%) were 2nd degree sprains, and 136 (6.5%) were 3rd degree sprains. Of the new
lateral ankle sprains, 1,957 (94%) had a documented date for return to play; 129 (6%) were censored cases of RTP. Since only new lateral ankle sprains were considered, the number of injured athletes and the number of injuries were equal. A majority of the injured athletes were male (64%) with 1st degree sprains (63%). The mean age of injured athletes was 16 (SD=1); approximately one quarter of injured athletes were in each of the high school grades (freshman, sophomore, junior, senior). Most injuries occurred in the regular season (75%) during competition (53%). A summary of demographic, injury, and competition characteristics is provided in Table 3.3.

Return to play probabilities, defined as the probability of returning to play during the specified time interval, given that RTP did not occur prior to that time interval, are presented by ankle sprain severity in Table 3.4. For 1st and 2nd degree lateral ankle sprains, the probability of RTP is highest 10-21 days after injury. For 3rd degree lateral ankle sprain, the probability of RTP is highest at least four weeks after injury. The log rank test identified a significant difference in survival curves between ankle sprain severity (p<.0001); there was a significant difference between 1st and 2nd degree survival curves (p<.0001) and 1st and 3rd degree survival curves (p<.0001), but there was no significant difference between 2nd and 3rd degree survival curves (p=.9). In the context of the RTP event measured in discrete time, failure probabilities are of more interest than survival probabilities. Failure for a specified time interval is the cumulative probability that RTP occurs prior to that time interval. The significant difference in survival curves is illustrated by plots of cumulative RTP probabilities in Figure 3.7.

The dependent variable of the discrete logistic regression model was coded so those who experienced the RTP event were compared to those who did not RTP (Table 3.5). There
was a significant interaction effect on RTP between time interval of return and ankle sprain severity ($p=.001$). Athletes who experienced a 1st degree sprain were 458% more likely to RTP in 1-2 days than athletes who experienced a 3rd degree sprain ($p=.001$), and 2nd degree sprains were 259% more likely to RTP in 1-2 days than 3rd degree sprains ($p=.02$). RTP in 3-6 days followed a similar trend; 1st degree sprains were 159% more likely ($p<.0001$) and 2nd degree sprains were 72% more likely ($p=.03$) to RTP than 3rd degree sprains. This trend continues through 7-9 day return and 10-21 day return. RTP comparisons between ankle sprain severities for more than 21 day return were not statistically significant. Time in season of injury had a significant effect on RTP. Athletes injured in the post-season were 42% less likely to return to play before the end of the season than athletes injured during the regular season ($p=.01$). Year in school, gender, height, weight, and exposure (competition vs practice) did not have a significant effect on return to play. The Hosmer-Lemeshow goodness of fit test was not significant ($\chi^2=5.6$, $p=.7$), suggesting that the model fit the data well.

### 3.5 Discussion

The literature on return to play has been largely descriptive in nature, and time to event analysis methods have not been heavily applied. While the failure probabilities presented in publications aimed at providing probability estimates of return to play do provide useful illustration of significant differences between survival curves between strata, these probabilities do not align with the mathematical probability of RTP. Kaplan-Meier survival estimates are not mathematically equivalent to RTP probabilities, requiring the use of the Life-Table method to estimate hazard probabilities. Further complicating the issue is the common practice of collecting T-RTP in discrete time intervals, which
requires time to event analysis techniques different than developing a Cox regression model. This paper provides detailed guidance to assist athletic trainers and sports medicine researchers with understanding the appropriate data structure required for time to event analysis and the methodology appropriate for analyzing discrete time RTP categories. Further, the programming statements and data example provide a detailed description of how these models are run and how they can be interpreted. The data example demonstrates how the life-table is useful for directly applicable information on expected time to return to play, while the discrete logistic regression model for this example highlights the relationship between severity and time to return to play.

3.5.1 Return to play probabilities by ankle sprain severity
Ankle sprain severity has a significant effect on estimated time to return to play. The more severe the sprain, the longer it takes to RTP. After an injury is diagnosed, an immediate discussion typically surrounds healing timelines and when the athlete will be cleared to return to participation. The RTP probabilities from the life-table can help athletes, coaches, and trainers determine reasonable expectations for return times based on sprain severity. For a 1\textsuperscript{st} degree sprain, athletes have roughly a 50\% chance of returning in about one week, but a much higher chance of returning in the third week after injury. For a 2\textsuperscript{nd} degree sprain, an athlete has the highest chance of RTP in the third week after injury. A 3\textsuperscript{rd} degree sprain would put an athlete returning at least four weeks after injury. Regardless of sprain severity, an athlete has a fairly high chance of RTP within one month. Equipping athletes with reasonable expectations regarding return times can allow them adequate time to heal without rushing to return. Further, availability of
RTP probabilities can ease coaching decisions concerning lineups without the injured player if an estimated return timeline is available.

3.5.2 Multivariate relationship between RTP, time and severity

Similar to the life-table finding that ankle sprain severity has an effect on T-RTP, the discrete logistic regression model indicates a statistically significant interaction between time and severity. The hazard odds of RTP within three weeks of injury are higher for 1st and 2nd degree sprains compared to 3rd degree sprains. This finding is consistent with more severe ankle sprains resulting in longer return times. Also of interest was the effect of gender and exposure on RTP. While only a moderately significant effect, male athletes were more likely to return during the season than females. Female athletes only comprised one-third of the sample, which could impact this result. Although studies have shown only minor gender differences in injury vulnerability [87, 88], there have been recent opinion reports that gender plays a role in injury frequency and healing timelines [45, 46]. Theories postulate that females are more vulnerable to injury, particularly ankle sprains, due to higher estrogen levels, less powerful muscles, increased likelihood of calcium and Vitamin D deficiency, and a wider pelvis resulting in a different alignment of the knee and ankle [45, 46]. Examining gender differences in RTP timelines would provide a significant contribution to the body of literature surrounding this issue. Injuries occurring during competition were marginally less likely to result in RTP during the season than injuries occurring during practice, perhaps indicating that competition injuries are more severe than practice injuries. Players injured during the post-season were significant less likely to RTP during the season than players injured during the regular season. This is likely due to the end of the study period occurring in such close
proximity to post-season play. Perhaps only injuries that occur in the regular season should be included in analysis to allow adequate time for follow up in future studies. While only high school athletes were considered in this analysis, the sample was equally representative of each of the four grade classifications. Year in school, and by extension, age, did not have an effect on RTP. These findings confirm that for athletes, coaches, and clinicians, ankle sprain severity should be the priority consideration when evaluating RTP timeline decisions.

3.6 Conclusion

Return to play has typically been summarized in a descriptive manner and excluded athletes with no follow up information regarding return. The application of time to event analysis methods for discrete time have been applied to RTP using ankle sprain injuries reported by high school athletes. These methods provide return time probabilities and explore multivariate relationships between RTP and demographic and competition characteristics while accounting for censored cases. The findings suggest that ankle sprain severity has the strongest impact on RTP timelines. Historically, RTP estimations have been based on anecdotal evidence and subjective opinion, but these can be inaccurate and, coupled with pressure from coaches, athletes, and other external influences, athletes can be cleared to RTP too early. Returning to play too soon can lead to recurrent injury and long-term conditions such as osteoarthritis. In practice, all evidence needs to be considered when making RTP decisions. The addition of evidence-based, objective indicators in the form of RTP probabilities to this body of evidence can provide more accurate estimates of return times and more confidence in returning an athlete to play.
Table 3.1 Distribution functions for discrete time to return to play

Let T be a discrete random variable that can take on values \( t_j, j = 1, 2, \ldots \), where \( t_1 < t_2 < \ldots \).

<table>
<thead>
<tr>
<th>Function</th>
<th>Formula</th>
<th>Definition</th>
</tr>
</thead>
<tbody>
<tr>
<td>Probability mass function (pmf)</td>
<td>( f(t_j) = P(TRTP = t_j) )</td>
<td>probability that RTP took place at time ( t_j )</td>
</tr>
<tr>
<td>Cumulative distribution function (cdf, or Failure)</td>
<td>( F(t) = P(TRTP &lt; t) )</td>
<td>probability that RTP occurred prior to time ( t )</td>
</tr>
<tr>
<td>Survival function</td>
<td>( S(t) = P(TRTP \geq t) )</td>
<td>probability that RTP occurred at time ( t ) or later</td>
</tr>
<tr>
<td>Hazard function</td>
<td>( h(t_j) = P(TRTP = t_j \mid TRTP \geq t_j) )</td>
<td>conditional probability of RTP at time ( t_j ) given the athlete has not returned to play prior to ( t_j )</td>
</tr>
<tr>
<td>Cumulative hazard function</td>
<td>( H(t) = \sum_{t_j \leq t} h(t_j) )</td>
<td>sum of hazard up to time ( t )</td>
</tr>
</tbody>
</table>

Table 3.2 Estimated season start and end dates for field and court sports

<table>
<thead>
<tr>
<th>Sport</th>
<th>Season start date</th>
<th>Season end date</th>
</tr>
</thead>
<tbody>
<tr>
<td>Football</td>
<td>August 1</td>
<td>December 1</td>
</tr>
<tr>
<td>Fall soccer (boys/girls)</td>
<td>August 1</td>
<td>November 15</td>
</tr>
<tr>
<td>Spring soccer (boys/girls)</td>
<td>December 15</td>
<td>April 15</td>
</tr>
<tr>
<td>Volleyball</td>
<td>August 1</td>
<td>November 15</td>
</tr>
<tr>
<td>Wrestling</td>
<td>November 1</td>
<td>March 15</td>
</tr>
<tr>
<td>Basketball (boys/girls)</td>
<td>November 1</td>
<td>March 15</td>
</tr>
<tr>
<td>Baseball/softball</td>
<td>February 1</td>
<td>June 1</td>
</tr>
</tbody>
</table>
Table 3.3 Demographic, injury, and competition characteristics for high school athletes with new lateral ankle sprain (N=2,086)

<table>
<thead>
<tr>
<th>Variable</th>
<th>n (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Gender</strong></td>
<td></td>
</tr>
<tr>
<td>Male</td>
<td>1337 (64.1)</td>
</tr>
<tr>
<td>Female</td>
<td>749 (35.9)</td>
</tr>
<tr>
<td><strong>Grade in school</strong></td>
<td></td>
</tr>
<tr>
<td>Freshman</td>
<td>450 (21.9)</td>
</tr>
<tr>
<td>Sophomore</td>
<td>524 (25.4)</td>
</tr>
<tr>
<td>Junior</td>
<td>535 (26.0)</td>
</tr>
<tr>
<td>Senior</td>
<td>550 (26.7)</td>
</tr>
<tr>
<td><strong>Sport</strong></td>
<td></td>
</tr>
<tr>
<td>Football</td>
<td>752 (36.1)</td>
</tr>
<tr>
<td>Boys fall soccer</td>
<td>127 (6.1)</td>
</tr>
<tr>
<td>Girls fall soccer</td>
<td>144 (6.9)</td>
</tr>
<tr>
<td>Volleyball</td>
<td>234 (11.2)</td>
</tr>
<tr>
<td>Boys basketball</td>
<td>342 (16.4)</td>
</tr>
<tr>
<td>Girls basketball</td>
<td>289 (13.8)</td>
</tr>
<tr>
<td>Wrestling</td>
<td>73 (3.5)</td>
</tr>
<tr>
<td>Baseball</td>
<td>43 (2.1)</td>
</tr>
<tr>
<td>Softball</td>
<td>82 (3.9)</td>
</tr>
<tr>
<td><strong>Ankle sprain severity</strong></td>
<td></td>
</tr>
<tr>
<td>1&lt;sup&gt;st&lt;/sup&gt; degree</td>
<td>1305 (62.5)</td>
</tr>
<tr>
<td>2&lt;sup&gt;nd&lt;/sup&gt; degree</td>
<td>645 (30.9)</td>
</tr>
<tr>
<td>3&lt;sup&gt;rd&lt;/sup&gt; degree</td>
<td>136 (6.5)</td>
</tr>
<tr>
<td><strong>Returned to play</strong></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>1957 (93.8)</td>
</tr>
<tr>
<td>No</td>
<td>129 (6.2)</td>
</tr>
<tr>
<td><strong>Time to return to play</strong></td>
<td></td>
</tr>
<tr>
<td>1-2 days</td>
<td>293 (14.0)</td>
</tr>
<tr>
<td>3-6 days</td>
<td>659 (31.6)</td>
</tr>
<tr>
<td>7-9 days</td>
<td>421 (20.2)</td>
</tr>
<tr>
<td>10-21 days</td>
<td>504 (24.2)</td>
</tr>
<tr>
<td>More than 21 days</td>
<td>209 (10.0)</td>
</tr>
<tr>
<td><strong>Time in season</strong></td>
<td></td>
</tr>
<tr>
<td>Preseason</td>
<td>458 (22.0)</td>
</tr>
<tr>
<td>Regular season</td>
<td>1561 (74.9)</td>
</tr>
<tr>
<td>Post season</td>
<td>64 (3.1)</td>
</tr>
</tbody>
</table>
Table 3.3 (continued) Demographic, injury, and competition characteristics for high school athletes with new lateral ankle sprain (N=2,086)

<table>
<thead>
<tr>
<th>Variable</th>
<th>n (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Type of exposure</strong></td>
<td></td>
</tr>
<tr>
<td>Competition</td>
<td>1113 (53.4)</td>
</tr>
<tr>
<td>Practice</td>
<td>973 (46.6)</td>
</tr>
<tr>
<td><strong>Competition site</strong></td>
<td></td>
</tr>
<tr>
<td>Home</td>
<td>139 (47.6)</td>
</tr>
<tr>
<td>Away</td>
<td>143 (49.0)</td>
</tr>
<tr>
<td>Neutral</td>
<td>10 (3.4)</td>
</tr>
<tr>
<td><strong>Competition time</strong></td>
<td></td>
</tr>
<tr>
<td>Warm-up</td>
<td>6 (2.1)</td>
</tr>
<tr>
<td>Beginning</td>
<td>47 (16.4)</td>
</tr>
<tr>
<td>Middle</td>
<td>159 (55.4)</td>
</tr>
<tr>
<td>End</td>
<td>73 (25.4)</td>
</tr>
<tr>
<td>Overtime</td>
<td>2 (0.7)</td>
</tr>
</tbody>
</table>

Table 3.4 Estimated return to play probabilities by ankle sprain severity

<table>
<thead>
<tr>
<th>Time interval</th>
<th>1st degree</th>
<th>2nd degree</th>
<th>3rd degree</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>P (95% CI)</td>
<td>P (95% CI)</td>
<td>P (95% CI)</td>
</tr>
<tr>
<td>1-2 days</td>
<td>.17 (.15, .19)</td>
<td>.10 (.08, .13)</td>
<td>.03 (.00, .06)</td>
</tr>
<tr>
<td>3-6 days</td>
<td>.42 (.39, .45)</td>
<td>.30 (.27, .34)</td>
<td>.21 (.14, .27)</td>
</tr>
<tr>
<td>7-9 days</td>
<td>.43 (.39, .47)</td>
<td>.32 (.27, .36)</td>
<td>.23 (.15, .31)</td>
</tr>
<tr>
<td>10-21 days</td>
<td>.71 (.66, .76)</td>
<td>.72 (.67, .78)</td>
<td>.53 (.42, .64)</td>
</tr>
<tr>
<td>More than 21 days</td>
<td>.56 (.44, .68)</td>
<td>.72 (.61, .84)</td>
<td>.79 (.64, .93)</td>
</tr>
</tbody>
</table>

*Probability of RTP during the time interval given that the athlete has not returned prior to that interval.
Table 3.5 Discrete time logistic regression of return to play on time, injury severity, athlete demographics, and competition characteristics for lateral ankle sprains in high school athletes \((n=5,132)\)

<table>
<thead>
<tr>
<th>Variable</th>
<th>Hazard odds ratio</th>
<th>95% Confidence interval for hazard odds ratio</th>
<th>(p)-value</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Year in school</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Freshman vs Senior</td>
<td>0.89</td>
<td>(0.74, 1.09)</td>
<td>.3</td>
</tr>
<tr>
<td>Sophomore vs Senior</td>
<td>0.87</td>
<td>(0.73, 1.04)</td>
<td>.1</td>
</tr>
<tr>
<td>Junior vs Senior</td>
<td>0.93</td>
<td>(0.78, 1.12)</td>
<td>.4</td>
</tr>
<tr>
<td><strong>Gender</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Male vs Female</td>
<td>1.21</td>
<td>(0.98, 1.49)</td>
<td>.1</td>
</tr>
<tr>
<td><strong>Height (inches)</strong></td>
<td>0.99</td>
<td>(0.97, 1.01)</td>
<td>.4</td>
</tr>
<tr>
<td><strong>Weight (inches)</strong></td>
<td>1.00</td>
<td>(0.99, 1.00)</td>
<td>.7</td>
</tr>
<tr>
<td><strong>Time in season</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Pre-season vs regular season</td>
<td>1.08</td>
<td>(0.91, 1.28)</td>
<td>.4</td>
</tr>
<tr>
<td>Post-season vs regular season</td>
<td>0.58</td>
<td>(0.40, 0.86)</td>
<td>.01</td>
</tr>
<tr>
<td><strong>Exposure</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Competition vs practice</td>
<td>0.87</td>
<td>(0.75, 1.00)</td>
<td>.1</td>
</tr>
<tr>
<td><strong>T-RTP * ankle sprain severity</strong></td>
<td></td>
<td></td>
<td>.001</td>
</tr>
<tr>
<td>1-2 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1(^{st}) degree vs 3(^{rd}) degree</td>
<td>5.58</td>
<td>(2.03, 15.31)</td>
<td>.001</td>
</tr>
<tr>
<td>2(^{nd}) degree vs 3(^{rd}) degree</td>
<td>3.59</td>
<td>(1.28, 10.08)</td>
<td>.02</td>
</tr>
<tr>
<td>3-6 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1(^{st}) degree vs 3(^{rd}) degree</td>
<td>2.59</td>
<td>(1.63, 4.12)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>2(^{nd}) degree vs 3(^{rd}) degree</td>
<td>1.72</td>
<td>(1.06, 2.79)</td>
<td>.03</td>
</tr>
<tr>
<td>7-9 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1(^{st}) degree vs 3(^{rd}) degree</td>
<td>2.45</td>
<td>(1.47, 4.07)</td>
<td>.001</td>
</tr>
<tr>
<td>2(^{nd}) degree vs 3(^{rd}) degree</td>
<td>1.49</td>
<td>(0.88, 2.55)</td>
<td>.1</td>
</tr>
<tr>
<td>10-21 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1(^{st}) degree vs 3(^{rd}) degree</td>
<td>2.11</td>
<td>(1.24, 3.60)</td>
<td>.01</td>
</tr>
<tr>
<td>2(^{nd}) degree vs 3(^{rd}) degree</td>
<td>2.23</td>
<td>(1.28, 3.88)</td>
<td>.005</td>
</tr>
<tr>
<td>more than 21 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1(^{st}) degree vs 3(^{rd}) degree</td>
<td>0.49</td>
<td>(0.21, 1.14)</td>
<td>.1</td>
</tr>
<tr>
<td>2(^{nd}) degree vs 3(^{rd}) degree</td>
<td>0.98</td>
<td>(0.40, 2.40)</td>
<td>.9</td>
</tr>
</tbody>
</table>

T-RTP=Time to return to play

Note: Five indicators for sport were included in the model.

Note: The main effects of both time interval and ankle sprain severity are not included in the table; they are included in a significant interaction and are not practically interpretable.
Figure 3.1 HS Rio\textsuperscript{TM} ankle sprain data (First 20 records out of 2086)

Figure 3.2 SAS Life-Table for 1\textsuperscript{st} degree ankle sprain in high school athletes
Figure 3.3 Life-Table dialogue box in SPSS

Figure 3.4 SPSS Life-Table for 1st degree ankle sprain in high school athletes
**Figure 3.5 Original ankle sprain data and pseudo-observations for discrete logistic regression model**

<table>
<thead>
<tr>
<th>obs</th>
<th>player_id</th>
<th>injury_date</th>
<th>season_end_date</th>
<th>TRTP</th>
<th>TRTP_c</th>
<th>event</th>
<th>ankle Sprain severity</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>59</td>
<td>09/09/2008</td>
<td>12/01/2008</td>
<td>1</td>
<td>returned to activity in 1-2 days</td>
<td>1</td>
<td>1</td>
</tr>
<tr>
<td>2</td>
<td>195</td>
<td>08/24/2007</td>
<td>12/01/2007</td>
<td>3</td>
<td>returned to activity in 7-9 days</td>
<td>1</td>
<td>1</td>
</tr>
<tr>
<td>3</td>
<td>274</td>
<td>08/09/2011</td>
<td>12/01/2011</td>
<td>3</td>
<td>returned to activity in 7-9 days</td>
<td>1</td>
<td>1</td>
</tr>
<tr>
<td>4</td>
<td>346</td>
<td>11/22/2008</td>
<td>12/01/2008</td>
<td>4</td>
<td>returned to activity in 10-21 days</td>
<td>1</td>
<td>1</td>
</tr>
<tr>
<td>5</td>
<td>375</td>
<td>11/13/2007</td>
<td>12/01/2007</td>
<td>4</td>
<td>returned to activity in 10-21 days</td>
<td>1</td>
<td>1</td>
</tr>
<tr>
<td>6</td>
<td>388</td>
<td>10/27/2010</td>
<td>12/01/2010</td>
<td>5</td>
<td>returned to activity in 12 days or more</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
<td>7</td>
<td>636</td>
<td>08/19/2009</td>
<td>12/01/2009</td>
<td>4</td>
<td>returned to activity in 10-21 days</td>
<td>1</td>
<td>2</td>
</tr>
<tr>
<td>8</td>
<td>727</td>
<td>11/24/2009</td>
<td>12/01/2009</td>
<td>3</td>
<td>returned to activity in 7-9 days</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
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<td>746</td>
<td>11/27/2010</td>
<td>12/01/2010</td>
<td>2</td>
<td>returned to activity in 7-9 days</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
<td>10</td>
<td>941</td>
<td>08/28/2009</td>
<td>11/15/2009</td>
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<td>returned to activity in 12 days or more</td>
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<td>1</td>
</tr>
<tr>
<td>11</td>
<td>1271</td>
<td>11/01/2010</td>
<td>11/15/2010</td>
<td>4</td>
<td>returned to activity in 10-21 days</td>
<td>0</td>
<td>3</td>
</tr>
<tr>
<td>12</td>
<td>1429</td>
<td>01/27/2011</td>
<td>03/15/2011</td>
<td>1</td>
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<td>3</td>
</tr>
<tr>
<td>13</td>
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<td>01/03/2011</td>
<td>03/15/2011</td>
<td>5</td>
<td>returned to activity in 12 days or more</td>
<td>1</td>
<td>2</td>
</tr>
<tr>
<td>14</td>
<td>1399</td>
<td>03/09/2012</td>
<td>03/15/2012</td>
<td>2</td>
<td>returned to activity in 3-6 days</td>
<td>0</td>
<td>3</td>
</tr>
<tr>
<td>15</td>
<td>1657</td>
<td>12/12/2008</td>
<td>03/15/2009</td>
<td>2</td>
<td>returned to activity in 3-6 days</td>
<td>1</td>
<td>2</td>
</tr>
<tr>
<td>16</td>
<td>1870</td>
<td>11/26/2011</td>
<td>03/15/2012</td>
<td>5</td>
<td>returned to activity in 22 days or more</td>
<td>0</td>
<td>3</td>
</tr>
<tr>
<td>17</td>
<td>1898</td>
<td>01/24/2008</td>
<td>03/15/2008</td>
<td>2</td>
<td>returned to activity in 1-2 days</td>
<td>1</td>
<td>3</td>
</tr>
<tr>
<td>18</td>
<td>2004</td>
<td>06/01/2011</td>
<td>06/01/2011</td>
<td>1</td>
<td>returned to activity in 1-2 days</td>
<td>0</td>
<td>2</td>
</tr>
<tr>
<td>19</td>
<td>2030</td>
<td>01/11/2009</td>
<td>06/01/2009</td>
<td>3</td>
<td>returned to activity in 1-2 days</td>
<td>1</td>
<td>3</td>
</tr>
<tr>
<td>20</td>
<td>2086</td>
<td>06/01/2011</td>
<td>06/01/2011</td>
<td>1</td>
<td>returned to activity in 1-2 days</td>
<td>0</td>
<td>1</td>
</tr>
</tbody>
</table>
Figure 3.6 Maximum likelihood estimates of Discrete-Time Logistic Model for HS Rio™ ankle sprain data
Figure 3.7 Cumulative RTP probabilities by lateral ankle sprain severity in high school athletes
4 Return to Play Probabilities after Knee Injury in High School Athletes

4.1 Introduction

Return to play (RTP) decisions have historically been determined using subjective reasoning; therefore, there is a need for more objective methods to assist in the determination of RTP to ensure athletes are not returning too soon. Typical research studies involving RTP report rates or proportions, but these measures can be inaccurate in instances where a lack of follow-up data excludes some athletes from analysis. Time to event analysis is a commonly used statistical analysis method that can provide more accurate estimates for time to return to play (T-RTP) by accounting for all injured athletes regardless of lack of follow up concerning RTP. These analysis techniques have not been heavily applied or related to RTP, and will provide athletic trainers, coaches, team physicians, and other members of the sports medicine team more accurate means to estimate return time. For prognosis, it is preferable to know the likelihood of an individual experiencing the outcome of interest within a given timeframe as opposed to summary results from cumulative risks and rates. In addition, it is likely that some athletes will not experience a return to play event, indicating the presence of censored cases. Time to event analysis can be applied to address both issues by generating evidence-based, objective estimates of when an athlete is likely to return following a given injury while accounting for censored cases. The accuracy of predicting T-RTP can be improved using these techniques, resulting in better patient care through education, improved coach-medical staff relations, and more efficient use of an athletic trainer’s clinical time.
Knee joint ligament sprains are the second most common lower extremity (LE) sports-related injury [2, 3, 65]. Medial collateral ligament (MCL) sprains are the most common ligament injuries at the knee and occur frequently in football [66], ice hockey [67], soccer [68], and skiing [69], all sports in which body movements causing high valgus stress at the knee are common [70, 71]. Incidence of anterior cruciate ligament (ACL) injury has recently risen significantly in adolescent athletes [7]. This rise could potentially be due to improvements in identification of injury through diagnostic testing [1], or an increase in participation in sports where the mechanisms of ACL injury are common. Basketball, football, and soccer participation has increased among adolescent athletes; these sports often require deceleration or change of direction forces [7]. Knee joint ligament sprains lead to time lost from activity, functional and chronic instability, and joint degeneration over time, even though surgery is not typical. For mild (Grade 1) MCL sprains, reported diminished functional capacity lasts for several weeks with RTP ranging from 4 to 19 days post-injury [75]; reported diminished functional capacity for moderate (Grade 2) MCL sprains ranges from 3 to 8 weeks [76]. ACL injury RTP guidelines suggest it could take between 4 and 8 weeks for full range of motion to return and swelling to subside [77]. As evidenced by these imprecise timeframes, healing and RTP timelines are difficult to predict based on tissue damage alone.

The extent of tissue damage and tissue healing timelines are not direct indicators of how long an athlete will be withheld from play, and many RTP decisions focus on resolution of symptoms. However, symptom resolution and restoration of function can follow a vastly different timeline than tissue healing, making a prognosis of time to RTP difficult. Current RTP estimations are predominantly based on subjective reports of the individual
patient’s symptoms and anecdotal, expert opinion of clinicians. While neither approach is
incorrect, there is limited clinical epidemiological evidence to substantiate current RTP
predictions for knee injuries. As a consequence, athletes may have unrealistic
expectations for RTP, contributing to poor compliance with rehabilitation programs,
disregarding symptoms, and pushing for earlier RTP, putting them at risk of subsequent
or increasingly severe injury. Recurrent knee injuries can lead to concomitant long-term
health issues; ligament damage to the knee can result in early-onset post-traumatic
osteoarthritis (OA), most often within 10 to 20 years after initial injury [9, 10]. For high
school athletes suffering ankle or knee injuries, this indicates the potential for
degenerative conditions before age 30. Posttraumatic OA symptoms include joint pain,
swelling, limited motion, and disability [11]; these symptoms are associated with
decreased physical activity, obesity, cardiovascular disease, and depression stemming
from loss of function and disability due to injury to the affected joint [12, 13]. To address
the potential significant public health issues related to knee joint ligament sprains in the
adolescent athlete population, there is a need for the development of effective strategies
to diminish the impact of joint injuries, improve compliance to rehabilitation programs,
and reduce the risk of reinjury, with the goal of avoiding long-term effects from injury
and the continued maintenance of joint health. It is critical that athletic trainers,
physicians, and coaches be able to accurately gauge when an athlete is safe to return to
play. The purpose of this study was to analyze RTP probabilities between different
single-ligament knee injuries using time to event analysis methodologies.
4.2 Methods

4.2.1 Design and sample

This observational study consisted of a secondary data analysis of athlete injury data from the High School RIO™ Injury Surveillance System (ISS) database. All US high schools with a National Athletic Trainers’ Association (NATA) certified athletic trainer (AT) were eligible for enrollment in this ISS; AT’s who enrolled their school completed the online “Exposure Report Form” for reportable injuries each week. New single-ligament knee sprains experienced by high school athletes from approximately 100 high schools in the US during regularly scheduled participation in school-sanctioned sports for seven academic years (2005-2006 through 2011-2012) were considered. Only field and court sports (football, boys/girls soccer, volleyball, wrestling, basketball, baseball, and softball) were included as these athletes were more likely to suffer knee injury due to the nature of the sport. The use of the HS RIO™ ISS database for the purpose of evaluating RTP probabilities was approved by the University Institutional Review Board.

Within the HS RIO™ ISS, a sprain was defined as injury to the ligamentous or capsular tissue [2, 83]. All knee joint ligament sprains that required the athlete to be removed from participation and diagnosed by the treating health care professional were reported, regardless of time lost from participation. No unique personal identifying information was contained in the dataset; the de-identification process of the data prohibited linking multiple injuries on the same athlete allowing only new injuries to be considered for analysis. A new injury was defined as a knee joint ligament sprain with an acute, traumatic onset of symptoms with no prior history of that injury. The ligaments under consideration include the anterior cruciate ligament (ACL), posterior cruciate ligament
(PCL), lateral collateral ligament (LCL), and medial collateral ligament (MCL) [89].
There were 1,049 documented new single-ligament knee injuries (classified under injury
type=“ligament sprain”) among field and court sport athletes with a known injury date.
RTP probabilities will be presented stratified by injured knee ligament.

4.2.2 Measures

Time to Return to Play Outcome. T-RTP was defined as the time between injury and
when the athlete was able to return to play. In the HS RIO™ ISS database, the number of
days the athlete was withheld from participation was collected in intervals of 1-2 days, 3-
6 days, 7-9 days, 10-21 days, and more than 21 days [83].

Censored Cases of Return to Play. Athletes who did not return to play after injury were
considered censored cases of RTP. A censored observation could have occurred in one of
six ways: 1) the athlete was determined to be medically disqualified for the season; 2) the
athlete was determined to be medically disqualified for their career; 3) the athlete chose
not to continue but was not medically disqualified; 4) the athlete was released from the
team but was not medically disqualified; 5) the athlete did not return for unspecified
reasons; or 6) the season ended before the athlete could return to play. In all cases,
censored observations were non-informative and right censored. To compute T-RTP
categories for censored cases, estimated season end dates for each sport were used as a
proxy for last date of contact. The number of days between injury date and season end
date was calculated, and each injury was then classified into the appropriate T-RTP
category. Estimated season start and end dates for each sport are listed in Table 4.1.
Athlete Demographics and Competition Characteristics. Demographic characteristics were documented for each injured athlete including year in school (freshman, sophomore, junior, senior), age in years, and gender. Competition characteristics at the time of injury were also documented, including sport in which the athlete was participating, time in season (preseason, regular season, postseason), type of exposure (competition or practice), competition site (home, away, neutral site), and competition time (warm-ups, beginning, middle, end, overtime). Indicators for sport (football, soccer, volleyball, wrestling, basketball, baseball/softball) were included in the model to account for differences in injury frequency among field and court sports that may influence RTP.

4.2.3 Data analysis

Data were summarized using descriptive statistics, including means and standard deviations or frequency distributions. The Life-Table survival analysis method was used to calculate return to play probability estimates stratified on injured knee ligament. The log-rank test was used to determine differences in RTP probabilities between strata; the Tukey adjustment for multiple comparisons was used for pairwise comparisons of RTP probabilities between strata. Discrete logistic regression was used to estimate hazard odds ratios for RTP due to the discrete nature of the T-RTP outcome; the Hosmer-Lemeshow test was used to assess the fit of the discrete logistic regression model. Data analysis was performed using SAS for Windows (version 9.4, SAS Institute, Inc., Cary NC, USA); an alpha level of .05 was used throughout.

4.3 Results

Of all new single-ligament knee injuries, 310 (29.5%) were ACL sprains, 27 (2.6%) were PCL sprains, 103 (9.8%) were LCL sprains, and 609 (58.1%) were MCL sprains. Of the
1,049 total knee joint ligament sprains, 721 (68.7%) had a documented date for return to play; 328 (31.3%) were censored cases of RTP. Since only new knee sprains were considered, the number of injured athletes and the number of injuries were equal. A summary of demographic, injury, and competition characteristics is provided in Table 4.2. A majority of the injured athletes were male (80.3%) with MCL sprains (58.1%). The mean age of injured athletes was 16 (SD=1); over half of the injured athletes were in their junior or senior years. Most injuries occurred in the regular season (76.9%) during competition (63.8%).

Return to play probabilities, defined as the probability of returning to play during the specified time interval, given that RTP did not occur prior to that time interval, are presented by injured knee ligament in Table 4.3. As expected, the probability of RTP for ACL injuries was low, even three weeks after injury. For PCL sprain, the probability of RTP increased from 0.2 at 3-6 days to 0.33 at 10-21 days to 0.5 four weeks after injury. The probability of RTP after LCL and MCL sprains is approximately 0.5 at 10-21 days after injury. The log rank test identified a significant difference in survival curves between injured knee ligaments (p<.0001); pairwise comparisons between ligaments resulted in significant differences for all comparisons (p<.0001 for all). In the context of the RTP event measured in discrete time, failure probabilities are of more interest than survival probabilities. Failure for a specified time interval is the cumulative probability that RTP occurs prior to that time interval. The significant difference in survival curves is illustrated by plots of cumulative RTP probabilities in Figure 4.1.

The dependent variable of the discrete logistic regression model was coded so those who experienced the RTP event were compared to those who did not RTP (Table 4.4). There
was a significant interaction effect on RTP between time interval of return and injured knee ligament ($p=.01$). Athletes who experienced ACL sprain were 78% less likely to RTP in 1-2 days than athletes with MCL sprain ($p=.002$), 81% less likely to return in 3-6 days ($p<.0001$), 91% less likely to return in 7-21 days ($p<.0001$), and 74% less likely to return 4 weeks after injury ($p<.0001$). Athletes who experienced LCL sprains were 213% more likely to return in 1-2 days than athletes with MCL sprain ($p=.02$), 73% more likely to return in 3-6 days ($p=.04$), and 103% more likely to return in 7-9 days ($p=.02$). PCL sprains were not included in the discrete logistic model due to the relatively small number (3%) of athletes who reported this injury; this small sample size led to inflated standard errors and unstable estimates. Type of exposure had a significant effect on RTP; athletes suffering a single-ligament knee sprain during competition were 25% less likely to RTP before the end of the season than athletes injured during practice ($p=.01$). Year in school, gender, and time in season did not have a significant effect on RTP. The Hosmer-Lemeshow goodness of fit test was not significant ($\chi^2=7.45$, $p=.5$), suggesting that the model fit the data well.

4.4 Discussion

As expected, MCL and ACL sprains were the most common occurring knee injuries among the high school athletes, and athletes who suffered an ACL injury had significantly lower RTP probabilities than other knee ligament injuries. After an injury is diagnosed, an immediate discussion typically surrounds healing timelines and when the athlete will be cleared to return to participation. The RTP probabilities from the life-table can help athletes, coaches, and athletic trainers determine reasonable expectations for return times based on injured knee ligament. Athletes injuring the ACL any time during
the season have only a 1 in 3 chance of returning before the end of the season. This finding is consistent with 4 to 8 week estimates from existing RTP guidelines given the study period defined by sport seasons [77]. Those suffering PCL sprain have a 50% chance of returning before season end. Athletes with LCL and MCL sprains have roughly a 50% chance of returning in the third week after injury, which is also consistent with existing RTP estimations [75]. Regardless of the knee ligament injured, athletes have a very small chance of RTP within two weeks of injury; RTP probabilities increase slightly for PCL, LCL and MCL injuries after two weeks. Equipping athletes with reasonable expectations regarding return times can allow them adequate time to heal without rushing to return. In addition to an individual athlete’s clinical symptoms, availability of RTP probabilities can ease coaching decisions concerning lineups without the injured player if an estimated return timeline is available.

Similar to the life-table finding that injured knee joint ligament has an effect on T-RTP, the discrete logistic regression model indicates a statistically significant interaction between time and injured ligament. The hazard odds of RTP for ACL sprain are significantly lower across all time points than MCL sprain. This finding is consistent with longer healing times for ACL injury than other knee ligaments [77]. Conversely, the hazard odds of RTP for LCL sprain are significantly higher the first week after injury than MCL sprain. Also of interest was the significant effect of exposure on RTP. Players injured during competition were less likely to return during the season than players injured in practice, suggesting that competition injuries are more severe and require longer healing times than practice injuries. No measure of injury severity for knee sprain was available in the HS RIO™ database, and would have the potential for unreliability as
this is a subjective measure when assessed immediately following injury by an athletic trainer. However, severity of injury should certainly be considered as a factor with the potential to influence T-RTP. Timing in season likely did not have an effect on RTP as most injuries across all ligaments occurred during the regular season. At least two-thirds of knee joint ligament sprains were suffered by male athletes, a result likely due to the high number of football players experiencing knee joint ligament sprains and consistent with the literature [7]. Interestingly, while female basketball players suffered more ACL and PCL injuries than male basketball players, gender did not have an effect on RTP. This result held even after limiting the model to ACL sprains in basketball players ($p=.2$).

This finding seems to suggest that although girls are more vulnerable to ACL injury [7, 40], RTP timelines are not different than boys. Further research is necessary to explore why girls are more susceptible to ACL injury and whether these risk factors have an effect on T-RTP. Regardless of slightly more knee sprains in junior and senior athletes, year in school, and by extension, age, did not have an effect on RTP.

Over 60% of knee ligament injuries were reported in football players. A sub-analysis of football injuries was conducted; a discrete logistic regression model was fit exploring the interaction between T-RTP and injured ligament controlling for year in school, time in season, and exposure (Table 5.4). The Hosmer-Lemeshow goodness of fit test was not significant ($\chi^2=7.41$, $p=.5$), suggesting that the model fit the data well. Findings were similar in direction of the hazard odds ratios and significance to findings across all sports with three exceptions. While year in school did not have an effect on RTP across all sports, freshman football players were significantly less likely to RTP before the end of the season after single-ligament knee sprain than senior football players ($p=.02$). Contrary
to findings across all sports, there was no significant difference in T-RTP between LCL and MCL sprains in the 3-6 day and 7-9 day return categories. The similarities in findings in this sub-analysis of football knee ligament injuries suggests that evaluating injuries and presenting RTP probabilities across all field and court sports is appropriate; however, sport-specific RTP probabilities may be of interest in certain applications. This should be considered when interpreting and generalizing results from this study.

4.4.1 Limitations

Researchers have only agreed upon two risk factors for knee ligament injury: being a female athlete and previous knee injury [79]. A significant limitation of this analysis was the inability to control for previous injury. Additionally, no measure of knee sprain severity was available. While it was expected for single-ligament PCL injuries to be rare, the relatively small number of injuries reported by the high school athletes does not allow for inclusion in the discrete logistic regression model. Further investigation of multi-ligament injuries is necessary to explore RTP probabilities after PCL injury. Further, due to the complex structure of the knee, multi-ligament injuries are more common than single ligament injuries [90]. If the LCL is injured, other structures in the joint are typically injured as well. Similarly, since the MCL helps protect the ACL against certain extreme knee forces, ACL injuries can occur alongside injuries to the MCL [78]. Multi-ligament injury analysis could provide sports medicine teams with more practically applicable information regarding T-RTP. Censored cases of RTP were more common for ACL and MCL sprains. The relatively large number of censored cases of ACL and MCL sprains in the higher time intervals, particularly in the more than 21 day return category, could potentially cause unstable estimates of RTP. This data collection issue highlights
the need for accurate RTP reporting. While the categorical collection of T-RTP on the Exposure Report Form is common in sports injury surveillance, it would be more efficient for investigation of RTP to collect exact return dates and allow for the precise calculation of T-RTP probability estimates using Cox regression models. This is particularly of concern for injuries that have healing timelines longer than three weeks, which is typically the largest time point in the categorical collection of RTP.

4.5 Conclusion

Return to play has typically been summarized in a descriptive manner and excluded athletes with no follow up information regarding return. The application of time to event analysis methods for discrete time have been applied to RTP using single-ligament knee sprain injuries reported by high school athletes. These methods provide return time probabilities and explore multivariate relationships between RTP and demographic and competition characteristics while accounting for censored cases. The findings suggest that injured knee ligament has a strong impact on RTP timelines. As expected, ACL injuries had the highest proportion of censored cases, and the least likelihood to RTP at all time points following injury. The likelihood of RTP following MCL or LCL injuries was far higher than that of the cruciate ligaments. Historically RTP estimations have been based on anecdotal evidence and subjective opinion, but these can be inaccurate and, coupled with pressure from coaches, athletes, and other external influences, athletes can be cleared to RTP too early. Returning to play too soon can lead to recurrent injury and the potential for long-term degenerative conditions. In practice, all evidence needs to be considered when making RTP decisions. The addition of evidence-based, objective indicators in the form of RTP probabilities to this body of evidence can provide the sports
medicine team with tools to determine more accurate estimates of return times and more confidence in returning an athlete to play.
<table>
<thead>
<tr>
<th>Sport</th>
<th>Season start date</th>
<th>Season end date</th>
</tr>
</thead>
<tbody>
<tr>
<td>Football</td>
<td>August 1</td>
<td>December 1</td>
</tr>
<tr>
<td>Fall soccer (boys/girls)</td>
<td>August 1</td>
<td>November 15</td>
</tr>
<tr>
<td>Spring soccer (boys/girls)</td>
<td>December 15</td>
<td>April 15</td>
</tr>
<tr>
<td>Volleyball</td>
<td>August 1</td>
<td>November 15</td>
</tr>
<tr>
<td>Wrestling</td>
<td>November 1</td>
<td>March 15</td>
</tr>
<tr>
<td>Basketball (boys/girls)</td>
<td>November 1</td>
<td>March 15</td>
</tr>
<tr>
<td>Baseball/softball</td>
<td>February 1</td>
<td>June 1</td>
</tr>
</tbody>
</table>
Table 4.2 Demographic, injury, and competition characteristics for high school athletes with new single-ligament knee sprain by injured ligament (N=1,049)

<table>
<thead>
<tr>
<th>Variable</th>
<th>ACL (n=310)</th>
<th>PCL (n=27)</th>
<th>LCL (n=103)</th>
<th>MCL (n=609)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>n (%)</td>
<td>n (%)</td>
<td>n (%)</td>
<td>n (%)</td>
</tr>
<tr>
<td>Gender</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Male</td>
<td>206 (66.5)</td>
<td>17 (63.0)</td>
<td>79 (76.7)</td>
<td>540 (88.7)</td>
</tr>
<tr>
<td>Female</td>
<td>104 (33.5)</td>
<td>10 (37.0)</td>
<td>24 (23.3)</td>
<td>69 (11.3)</td>
</tr>
<tr>
<td>Grade in school</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Freshman</td>
<td>48 (15.6)</td>
<td>3 (11.1)</td>
<td>27 (26.5)</td>
<td>110 (18.2)</td>
</tr>
<tr>
<td>Sophomore</td>
<td>63 (20.4)</td>
<td>4 (14.8)</td>
<td>17 (16.7)</td>
<td>144 (23.8)</td>
</tr>
<tr>
<td>Junior</td>
<td>109 (35.4)</td>
<td>11 (40.8)</td>
<td>24 (23.5)</td>
<td>160 (26.5)</td>
</tr>
<tr>
<td>Senior</td>
<td>88 (28.6)</td>
<td>9 (33.3)</td>
<td>34 (33.3)</td>
<td>190 (31.5)</td>
</tr>
<tr>
<td>Sport</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Football</td>
<td>156 (50.3)</td>
<td>10 (37.1)</td>
<td>55 (53.4)</td>
<td>420 (69.0)</td>
</tr>
<tr>
<td>Boys fall soccer</td>
<td>19 (6.1)</td>
<td>2 (7.4)</td>
<td>9 (8.8)</td>
<td>26 (4.3)</td>
</tr>
<tr>
<td>Girls fall soccer</td>
<td>24 (7.7)</td>
<td>2 (7.4)</td>
<td>7 (6.8)</td>
<td>25 (4.1)</td>
</tr>
<tr>
<td>Volleyball</td>
<td>15 (4.8)</td>
<td>0 (0.0)</td>
<td>7 (6.8)</td>
<td>5 (0.8)</td>
</tr>
<tr>
<td>Boys basketball</td>
<td>16 (5.2)</td>
<td>1 (3.7)</td>
<td>2 (1.9)</td>
<td>28 (4.6)</td>
</tr>
<tr>
<td>Girls basketball</td>
<td>53 (17.1)</td>
<td>5 (18.5)</td>
<td>10(9.7)</td>
<td>24 (3.9)</td>
</tr>
<tr>
<td>Wrestling</td>
<td>12 (3.9)</td>
<td>3 (11.1)</td>
<td>13 (12.6)</td>
<td>59 (9.7)</td>
</tr>
<tr>
<td>Baseball</td>
<td>3 (1.0)</td>
<td>1 (3.7)</td>
<td>0 (0.0)</td>
<td>7 (1.1)</td>
</tr>
<tr>
<td>Softball</td>
<td>12 (3.9)</td>
<td>3 (11.1)</td>
<td>0 (0.0)</td>
<td>15 (2.5)</td>
</tr>
<tr>
<td>Returned to play</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>97 (31.3)</td>
<td>19 (70.4)</td>
<td>91 (88.4)</td>
<td>514 (84.4)</td>
</tr>
<tr>
<td>No</td>
<td>213 (68.7)</td>
<td>8 (29.6)</td>
<td>12 (11.6)</td>
<td>95 (15.6)</td>
</tr>
<tr>
<td>Time to return to play</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1-2 days</td>
<td>6 (1.9)</td>
<td>2 (7.4)</td>
<td>14 (13.6)</td>
<td>46 (7.6)</td>
</tr>
<tr>
<td>3-6 days</td>
<td>16 (5.2)</td>
<td>5 (18.5)</td>
<td>28 (27.1)</td>
<td>127 (20.8)</td>
</tr>
<tr>
<td>7-9 days</td>
<td>7 (2.3)</td>
<td>2 (7.4)</td>
<td>19 (18.5)</td>
<td>85 (13.9)</td>
</tr>
<tr>
<td>10-21 days</td>
<td>25 (8.0)</td>
<td>6 (22.2)</td>
<td>23 (22.3)</td>
<td>182 (30.0)</td>
</tr>
<tr>
<td>More than 21 days</td>
<td>256 (82.6)</td>
<td>12 (44.5)</td>
<td>19 (18.5)</td>
<td>169 (27.7)</td>
</tr>
<tr>
<td>Time in season</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Preseason</td>
<td>57 (18.4)</td>
<td>3 (11.1)</td>
<td>24 (23.3)</td>
<td>117 (19.2)</td>
</tr>
<tr>
<td>Regular season</td>
<td>239 (77.1)</td>
<td>22 (81.5)</td>
<td>75 (72.8)</td>
<td>471 (77.3)</td>
</tr>
<tr>
<td>Post season</td>
<td>14 (4.5)</td>
<td>2 (7.4)</td>
<td>4 (3.9)</td>
<td>21 (3.5)</td>
</tr>
</tbody>
</table>

ACL=Anterior cruciate ligament; PCL=Posterior cruciate ligament; LCL=lateral collateral ligament; MCL=medial collateral ligament
Table 4.2 (continued) Demographic, injury, and competition characteristics for high school athletes with new single-ligament knee sprain by injured ligament (N=1,049)

<table>
<thead>
<tr>
<th>Variable</th>
<th>ACL (n=310)</th>
<th>PCL (n=27)</th>
<th>LCL (n=103)</th>
<th>MCL (n=609)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>n (%)</td>
<td>n (%)</td>
<td>n (%)</td>
<td>n (%)</td>
</tr>
<tr>
<td>Type of exposure</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Competition</td>
<td>222 (71.6)</td>
<td>15 (55.6)</td>
<td>64 (62.1)</td>
<td>368 (60.4)</td>
</tr>
<tr>
<td>Practice</td>
<td>88 (28.4)</td>
<td>12 (44.4)</td>
<td>39 (37.9)</td>
<td>241 (39.6)</td>
</tr>
<tr>
<td>Competition site</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Home</td>
<td>33 (61.1)</td>
<td>0 (0.0)</td>
<td>3 (27.3)</td>
<td>37 (43.5)</td>
</tr>
<tr>
<td>Away</td>
<td>20 (37.0)</td>
<td>3 (100.0)</td>
<td>8 (72.7)</td>
<td>41 (48.2)</td>
</tr>
<tr>
<td>Neutral</td>
<td>1 (1.9)</td>
<td>0 (0.0)</td>
<td>0 (0.0)</td>
<td>7 (8.3)</td>
</tr>
<tr>
<td>Competition time</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Warm-up</td>
<td>2 (3.8)</td>
<td>0 (0.0)</td>
<td>0 (0.0)</td>
<td>0 (0.0)</td>
</tr>
<tr>
<td>Beginning</td>
<td>7 (13.2)</td>
<td>0 (0.0)</td>
<td>0 (0.0)</td>
<td>11 (12.9)</td>
</tr>
<tr>
<td>Middle</td>
<td>29 (54.7)</td>
<td>3 (100.0)</td>
<td>7 (70.0)</td>
<td>43 (50.6)</td>
</tr>
<tr>
<td>End</td>
<td>15 (28.3)</td>
<td>0 (0.0)</td>
<td>3 (30.0)</td>
<td>31 (36.5)</td>
</tr>
<tr>
<td>Overtime</td>
<td>0 (0.0)</td>
<td>0 (0.0)</td>
<td>0 (0.0)</td>
<td>0 (0.0)</td>
</tr>
</tbody>
</table>

ACL=Anterior cruciate ligament; PCL=Posterior cruciate ligament; LCL=lateral collateral ligament; MCL=medial collateral ligament
Table 4.3 Estimated return to play probabilities by injured knee ligament

<table>
<thead>
<tr>
<th>Time interval</th>
<th>ACL</th>
<th>PCL</th>
<th>LCL</th>
<th>MCL</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>$P_j$ (95% CI)</td>
<td>$P_j$ (95% CI)</td>
<td>$P_j$ (95% CI)</td>
<td>$P_j$ (95% CI)</td>
</tr>
<tr>
<td>1-2 days</td>
<td>.02 (.00, .03)</td>
<td>.07 (.00, .17)</td>
<td>.14 (.07, .20)</td>
<td>.07 (.05, .09)</td>
</tr>
<tr>
<td>3-6 days</td>
<td>.05 (.03, .07)</td>
<td>.20 (.04, .36)</td>
<td>.31 (.22, .41)</td>
<td>.23 (.19, .26)</td>
</tr>
<tr>
<td>7-9 days</td>
<td>.02 (.00, .04)</td>
<td>.10 (.00, .23)</td>
<td>.31 (.20, .43)</td>
<td>.20 (.16, .23)</td>
</tr>
<tr>
<td>10-21 days</td>
<td>.08 (.05, .11)</td>
<td>.33 (.12, .55)</td>
<td>.53 (.38, .68)</td>
<td>.50 (.45, .56)</td>
</tr>
<tr>
<td>More than 21 days</td>
<td>.32 (.25, .40)</td>
<td>.50 (.15, .85)</td>
<td>.59 (.33, .85)</td>
<td>.65 (.57, .74)</td>
</tr>
</tbody>
</table>

$^1$Probability of RTP during the time interval given that the athlete has not returned prior to that interval.

ACL=Anterior cruciate ligament; PCL=Posterior cruciate ligament; LCL=lateral collateral ligament; MCL=medial collateral ligament
Table 4.4 Discrete time logistic regression of return to play on time, injured knee ligament, athlete demographics, and competition characteristics for new single-ligament knee sprains in high school athletes (n=3,852)

<table>
<thead>
<tr>
<th>Variable</th>
<th>Hazard odds ratio</th>
<th>95% Confidence interval for hazard odds ratio</th>
<th>p-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Year in school</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Freshman vs Senior</td>
<td>0.83</td>
<td>(0.63, 1.09)</td>
<td>.2</td>
</tr>
<tr>
<td>Sophomore vs Senior</td>
<td>1.07</td>
<td>(0.83, 1.37)</td>
<td>.6</td>
</tr>
<tr>
<td>Junior vs Senior</td>
<td>1.02</td>
<td>(0.80, 1.28)</td>
<td>.9</td>
</tr>
<tr>
<td>Gender</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Male vs Female</td>
<td>1.01</td>
<td>(0.69, 1.49)</td>
<td>.9</td>
</tr>
<tr>
<td>Time in season</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Pre-season vs regular season</td>
<td>1.05</td>
<td>(0.81, 1.36)</td>
<td>.7</td>
</tr>
<tr>
<td>Post-season vs regular season</td>
<td>0.69</td>
<td>(0.41, 1.16)</td>
<td>.2</td>
</tr>
<tr>
<td>Exposure</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Competition vs practice</td>
<td>0.75</td>
<td>(0.60, 0.93)</td>
<td>.01</td>
</tr>
<tr>
<td>T-RTP * injured knee ligament</td>
<td></td>
<td></td>
<td>.01</td>
</tr>
<tr>
<td>1-2 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.22</td>
<td>(0.09, 0.57)</td>
<td>.002</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>2.13</td>
<td>(1.12, 4.06)</td>
<td>.02</td>
</tr>
<tr>
<td>3-6 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.19</td>
<td>(0.11, 0.33)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>1.73</td>
<td>(1.05, 2.84)</td>
<td>.03</td>
</tr>
<tr>
<td>7-9 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.09</td>
<td>(0.04, 0.21)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>2.03</td>
<td>(1.12, 3.70)</td>
<td>.02</td>
</tr>
<tr>
<td>10-21 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.09</td>
<td>(0.05, 0.14)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>1.10</td>
<td>(0.57, 2.12)</td>
<td>.8</td>
</tr>
<tr>
<td>more than 21 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.26</td>
<td>(0.17, 0.41)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>0.82</td>
<td>(0.31, 2.15)</td>
<td>.7</td>
</tr>
</tbody>
</table>

T-RTP=Time to return to play; ACL=Anterior cruciate ligament; PCL=Posterior cruciate ligament; LCL=lateral collateral ligament; MCL=medial collateral ligament

Note: Five indicators for sport were included in the model.
Note: PCL sprains were not included in the model; the relatively small sample size (n=27) resulted in inflated standard errors and unstable estimates.
Note: The main effects of both time interval and injured knee ligament are not included in the table; they are included in a significant interaction and are not practically interpretable.
Table 4.5 Discrete time logistic regression of return to play on time, injured knee ligament, athlete demographics, and competition characteristics for new single-ligament knee sprains in high school football athletes ($n=2,367$)

<table>
<thead>
<tr>
<th>Variable</th>
<th>Hazard odds ratio</th>
<th>95% Confidence interval for hazard odds ratio</th>
<th>$p$-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Year in school</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Freshman vs Senior</td>
<td>0.67</td>
<td>(0.47, 0.94)</td>
<td>.02</td>
</tr>
<tr>
<td>Sophomore vs Senior</td>
<td>0.93</td>
<td>(0.68, 1.26)</td>
<td>.6</td>
</tr>
<tr>
<td>Junior vs Senior</td>
<td>0.90</td>
<td>(0.67, 1.21)</td>
<td>.5</td>
</tr>
<tr>
<td>Time in season</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Pre-season vs regular season</td>
<td>1.15</td>
<td>(0.84, 1.58)</td>
<td>.4</td>
</tr>
<tr>
<td>Post-season vs regular season</td>
<td>0.68</td>
<td>(0.35, 1.32)</td>
<td>.3</td>
</tr>
<tr>
<td>Exposure</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Competition vs practice</td>
<td>0.70</td>
<td>(0.53, 0.91)</td>
<td>.01</td>
</tr>
<tr>
<td>T-RTP* injured knee ligament</td>
<td></td>
<td></td>
<td>.01</td>
</tr>
<tr>
<td>1-2 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.17</td>
<td>(0.04, 0.72)</td>
<td>.02</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>2.93</td>
<td>(1.34, 6.43)</td>
<td>.01</td>
</tr>
<tr>
<td>3-6 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.24</td>
<td>(0.12, 0.46)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>1.63</td>
<td>(0.83, 3.19)</td>
<td>.2</td>
</tr>
<tr>
<td>7-9 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.07</td>
<td>(0.02, 0.28)</td>
<td>.0002</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>1.58</td>
<td>(0.63, 3.94)</td>
<td>.3</td>
</tr>
<tr>
<td>10-21 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.07</td>
<td>(0.03, 0.14)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>1.13</td>
<td>(0.46, 2.77)</td>
<td>.8</td>
</tr>
<tr>
<td>more than 21 day return</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ACL vs MCL</td>
<td>0.23</td>
<td>(0.13, 0.42)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>LCL vs MCL</td>
<td>0.88</td>
<td>(0.22, 3.49)</td>
<td>.9</td>
</tr>
</tbody>
</table>

*Time to return to play

ACL=Anterior cruciate ligament; PCL=Posterior cruciate ligament; LCL=lateral collateral ligament; MCL=medial collateral ligament

Note: PCL sprains were not included in the model; the relatively small sample size ($n=10$) resulted in inflated standard errors and unstable estimates.

Note: The main effects of both time interval and injured knee ligament are not included in the table; they are included in a significant interaction and are not practically interpretable.
Figure 4.1 Cumulative RTP probabilities by sprained knee ligament in high school athletes
5 Discussion and Conclusions

5.1 Summary

With approximately 30 million youth between the ages of 5 and 18 years in the US participating in organized sports [7], there is high potential for young athletes to suffer acute sports-related injuries or overuse injuries. It is estimated that 38% of high school athletes will suffer sports-related injuries requiring treatment by a physician [7]; however, actual percentages may be higher due to underreporting or failure to seek treatment [14]. Ankle ligament sprains are the most common injury across field and court sports [1, 8], and incidence rates have been shown to be higher in adolescents than adults [51]. Knee sprains are the second most common LE sports-related injury [2, 3, 65]. MCL sprains are the most common ligament injury at the knee, and ACL injuries have recently risen significantly in college and adolescent athletes [1, 7, 70, 71].

After injury, the athlete, coaches, and sports medicine team work together to formulate a RTP plan. There is a high degree of variability among clinicians regarding factors considered for RTP [15], resulting in a call to develop a more objective method to determine when it is safe for an athlete to resume participation after injury. A three-step decision making process has been proposed that recommends the evaluation of health status, participation risk, and consideration of the influence of outside factors surrounding the athlete [16]. The validation of this process illustrates the need for more objective methods to evaluate the risk of return to play after injury. Extensive documentation of risk factors for sports-related lateral ankle sprain and knee ligament injury exists in the literature; however, there is little agreement among researchers on factors that affect
RTP. One thing that has been agreed upon, however, is that premature RTP is a predisposing factor for subsequent injury [8]. Recurrent knee and ankle injuries can lead to concomitant long-term degenerative joint diseases. Ligament damage to the ankle or knee could potentially result in early-onset, post-traumatic OA, most often within 10 to 20 years after initial injury [9, 10]. For high school athletes suffering ankle or knee injuries, this indicates they could suffer degenerative conditions before the age of 30. Posttraumatic OA is typically reported in younger patients, and is significantly associated with decreased physical activity, obesity, cardiovascular disease, and depression stemming from loss of function and disability due to injury to the affected joint [12, 13]. This potentially significant public health issue related to ankle and knee sprains necessitates the need for development of effective strategies to diminish the impact of joint injuries, improve compliance with rehabilitation programs, and reduce the risk of subsequent injury, with the goal of avoiding long-term effects from injury and the continued maintenance of joint health.

Time to event analysis is a method of generating evidence-based, objective estimates of when an athlete is likely to RTP following a given injury. The accuracy of predicting T-RTP after ankle and knee sprains can be improved using time to event techniques, resulting in better patient care through education, improved coach-medical staff relations, and more efficient use of athletic trainer’s clinical time. Further, these analyses can improve clinical decision-making by adding research-based evidence and also incorporating the prognosis aspect to educate the patient and coaches about return to play timelines. Detailed guidance has been provided to assist athletic trainers and sports medicine researchers with understanding the appropriate data structure and programming
statements required for T-RTP analysis and the methodology appropriate for analyzing discrete time RTP categories.

For 1st and 2nd degree lateral ankle sprains, the probability of RTP is highest 10-21 days after injury. For 3rd degree lateral ankle sprain, the probability of RTP is highest at least four weeks after injury. There was a significant interaction effect on RTP between time interval of return and ankle sprain severity \((p=.001)\). Athletes who experienced a 1st degree sprain were 458% more likely to RTP in 1-2 days than athletes who experienced a 3rd degree sprain \((p=.001)\), and 2nd degree sprains were 259% more likely to RTP in 1-2 days than 3rd degree sprains \((p=.02)\). RTP in 3-6 days followed a similar trend; 1st degree sprains were 159% more likely \((p<.0001)\) and 2nd degree sprains were 72% more likely \((p=.03)\) to RTP than 3rd degree sprains. This trend continues through 7-9 day return and 10-21 day return. RTP comparisons between ankle sprain severities for more than 21 day return were not statistically significant. Time in season had a significant effect on RTP; athletes injured in the post-season were 42% less likely to return to play before the end of the season than athletes injured during the regular season \((p=.01)\). Gender had a marginal effect on RTP \((p=.1)\); male athletes were 21% more likely to return to play than female athletes. Year in school, height, weight, and exposure (competition vs practice) did not have a significant effect on return to play.

As expected, the probability of RTP for ACL injuries was low, even three weeks after injury. For PCL sprain, the probability of RTP increased from 0.2 at 3-6 days to 0.33 at 10-21 days to 0.5 four weeks after injury. The probability of RTP after LCL and MCL sprains is approximately 0.5 at 10-21 days after injury. There was a significant interaction effect on RTP between time interval of return and injured knee ligament \((p=.01)\). Athletes
who experienced ACL sprain were 78% less likely to RTP in 1-2 days than athletes with MCL sprain \((p=.002)\), 81% less likely to return in 3-6 days \((p<.0001)\), 91% less likely to return in 7-21 days \((p<.0001)\), and 74% less likely to return 4 weeks after injury \((p<.0001)\). Athletes who experienced LCL sprains were 213% more likely to return in 1-2 days than athletes with MCL sprain \((p=.02)\), 73% more likely to return in 3-6 days \((p=.04)\), and 103% more likely to return in 7-9 days \((p=.02)\). PCL sprains were not included in the discrete logistic model due to the relatively small number (3%) of athletes who reported this injury; this small sample size led to inflated standard errors and unstable estimates. Type of exposure had a significant effect on RTP; athletes suffering a single-ligament knee sprain during competition were 25% less likely to RTP before the end of the season than athletes injured during practice \((p=.01)\). Year in school, gender, and time in season did not have a significant effect on RTP. In a subset analysis of ACL injuries in basketball players, gender still had no significant effect on RTP \((p=.2)\).

The findings in this study provide a valuable start to addressing one component of the incredibly complex RTP question. Incorporating these RTP probability estimates into the participation risk component of a decision-based RTP model will provide stronger guidance for RTP decision, leading to better patient care, a reduction in reinjury rates, and a decrease in prevalence of early-onset OA among athletes.

5.2 Implications

The ability to return to play has long been a pivotal question posed by athletes and coaches after injury, and few prognostic indicators have been explored to indicate when RTP will occur. Clinicians have relied on their personal experience to predict when an athlete might return, and RTP decisions vary tremendously between clinicians [15].
Coaches are left with little information to formulate coaching decisions such as line-ups and substitutions. Further, this can potentially contribute to poor compliance with rehabilitation programs due to athletes pushing for earlier RTP, which can lead to subsequent or increasingly severe injuries [8]. A further consequence is the potential for concomitant long-term health issues such as OA [9, 91]. The estimates of T-RTP probabilities presented here contribute to the development of evidence-based, objective strategies to determine when it is safe for an athlete to RTP. Avoiding premature RTP can help diminish the impact of joint injuries, improve compliance with rehabilitation programs, and reduce the risk of reinjury, with the goal of avoiding long-term effects from injury and the continued maintenance of joint health.

5.3 Strengths and limitations

Alternative methods of reporting prognostic information exist, but a lack of certain elements limit their application. The advantage of time to event analyses is through the inclusion of censored cases, in other words, the inclusion within the analysis of those who are still waiting to experience the event of interest during the study timeframe. Specific to RTP, these analyses include data on athletes who do not return before the season is over, thereby limiting the bias associated with these analyses. In contrast, the reporting of proportions (e.g., 8/10 athletes returned after 3 days) eliminates data by not including those who did not return. Similarly, reporting the average timeframe for return also eliminates censored cases and may provide misleading information. This underestimation is illustrated with lateral ankle sprain data in high school athletes in Table 5.1. The estimated epidemiological incidence proportion (IP) is defined as the proportion of injured athletes that RTP during a specified time interval out of the number of injured
athletes that have not yet returned to play at the beginning of the interval [27]. When the number of censored cases is small, the IP and Lifetable RTP probabilities are nearly identical. However, as the number of censored cases increases, the IP is smaller than the Lifetable RTP probabilities. For 1st degree LAS, 5% of injuries were censored. Excluding censored cases, it is estimated that 33% of athletes will return 22 or more days after injury; however, accounting for censored cases, athletes with 1st degree LAS have an estimated 56% chance of return 22 or more days after injury. This underestimation is clear in LAS where only 6% of cases are censored; the differences would be even greater in the presence of a higher proportion of censored cases. More than 30% of single-ligament knee sprains in high school athletes were censored. Reporting proportions of RTP for knee sprains will significantly underestimate T-RTP; probability estimates determined using time to event analysis methods will be much more accurate in estimating T-RTP for knee ligament injuries. The literature on return to play has been largely descriptive in nature, and time to event analysis methods have only recently been applied to T-RTP [32, 92]. While the failure probabilities based on Kaplan-Meier survival estimates presented in these publications do provide useful illustration of significant differences between RTP survival curves between strata, these probabilities are not equivalent to the mathematical probability of RTP. The Life-Table method detailed and utilized in this study is more appropriate in this setting as it allows the direct computation of hazard probability estimates, which are equivalent to the probability of RTP during a specified time interval given the injured athlete has not returned prior to the specified time interval [31].
Participation in certain sports has been identified as a risk factor for injury [7, 8, 51, 61, 66-69]. To control for the observed differences in injury risk between types of sports, the decision was made to restrict analyses to only field and court sports in this study. Further, an indicator for sport was included in the discrete logistic regression models to control for differences in RTP between sports. Football and basketball dominated ankle and knee injury rates; girls basketball had higher incidence rates for single-ligament knee sprains than boys basketball. Sub-analyses were conducted to determine whether it was appropriate to report RTP probabilities and hazard odds ratios across all field and court sports. Findings from these sub-analyses did not indicate that T-RTP results were different enough to warrant stratification on sport at this time. There are certain applications in which sport-specific RTP probabilities may be of interest; this should be considered when interpreting and generalizing results from this study.

The HS RIO™ ISS has been established as a tracking system to determine incidence rates of injuries in the high school athlete population, as well as provide a mechanism by which to investigate risk factors for injury with the goal of reducing injury rates in high school athletes to the lowest possible level while still encouraging participation in organized sports [83]. NATA certified ATs volunteer to participate as reporters, and are then categorized into 8 sampling strata by geographic location and high school size. Schools from each substrata are randomly selected until a sample of 100 schools are identified. If an AT drops from participation, their school will be replaced by another school within the substrata. Participating ATs are offered $300-$400 depending on the number of sports in which they report data and individualized injury reports following the study’s conclusion [83]. While this sampling methodology is useful in maintaining a
nationally representative sample, the incentive to participate has the potential to bias results based on these data. Of particular concern is the potential for surveillance bias, which can result in an overestimation of the number of injuries in the ISS [93]. ATs may be inclined to identify an injury that would otherwise not be reported, particularly those with mild severity.

While the HS RIO™ ISS is useful for identifying risk factors for injury, documentation is lacking for factors that may influence RTP. Patient demographics, symptoms, and clinical injury evaluation information are well documented, but participation risk and decision modifiers identified for consideration in the three-step decision based RTP model are not collected on the exposure report form [16, 83]. In the analyses presented here, prior injury information was not available for examination. A proxy for injury severity for lateral ankle sprain was incorporated into the analyses; however, this was not feasible for knee ligament injuries. Information on injury protection (i.e., use of taping or bracing) after diagnosis was not reported. It is assumed that injuries reported to the HS RIO™ ISS are not verified after diagnostic testing or more confirmatory diagnoses methods by a clinician. This could potentially bias injury rates to ankle and knee ligaments as specific injured ligaments could be missed by ATs on the field, or ligaments could be identified as injured when they weren’t. Competition site (home, away, neutral site) and competition time (warm-ups, beginning, middle, end, overtime) were not included in analyses due to an abundance of missing data (86% missing for both). These competition characteristics were not identified as potential factors affecting T-RTP, so their exclusion is not of particular concern. However, it would be an addition to the body of knowledge surrounding T-RTP to explore the effect of these characteristics further.
More accurate reporting of competition site and time is suggested for future injury surveillance. Missing data for other study variables did not exceed 10%, eliminating the need to account for missing data in analyses.

A significant limitation to the use of HS RIO™ data for evaluating T-RTP after ankle and knee injury is that it does not capture non-time-loss (NTL) injuries aside from fractures, concussions, heat-related injuries, and dental injuries [83, 94]. Mild ankle and knee sprains could be classified as NTL injuries, and are missed in analyses conducted using this database. The National Athletic Treatment, Injury and Outcomes Network (NATION) ISS was developed partially in response to concerns about underreporting of NTL injuries [94]. The use of data from this ISS would account for NTL ankle and knee injuries, providing more accurate and robust estimates of RTP probabilities.

Another significant issue in the HS RIO™ ISS is the common practice of collecting T-RTP in discrete time intervals [83], which requires the use of the discrete logistic regression model to generate hazard odds ratios [31]. Censored cases of RTP were far more common for knee sprains of the ACL and MCL than LCL and PCL, particularly in the more than 21 day return category. The relatively large number of censored cases could potentially cause unstable estimates of RTP in those time intervals. In addition, the convention adopted in this study was to estimate T-RTP for censored cases by calculating the time between injury and end of season. This method does not provide adequate follow-up to ensure enough time for athletes to RTP, particularly those injured closer to the end of the season. This data collection issue highlights the need for accurate RTP reporting. While the categorical collection of T-RTP on the High School RIO™ ISS Exposure Report Form is common in sports injury surveillance, it would be more
efficient for investigation of RTP to collect exact return dates for a follow-up period through the beginning of the next season and allow for the precise calculation of T-RTP probability estimates using Cox regression models. In addition to more accurate estimates, statistical power to detect group differences is higher with a continuous outcome [95]. This is of particular concern for injuries that have healing timelines longer than three weeks, which is typically the largest time point in the categorical collection of RTP.

5.4 Future research

5.4.1 Parametric Accelerated Failure Time model

If the survival distribution is known, parametric methods could be employed to evaluate T-RTP [28, 31]. These parametric models accommodate all types of censored data, as well as allowing the ability to test certain hypotheses about the shape of the hazard function. A Cox model only gives the nonparametric estimates of the survival function, which can sometimes be difficult to interpret. The parametric method for investigating multiple variable relationships is the Accelerated Failure Time (AFT) model. Under this framework, the variance of \( \varepsilon_i \) varies from one data set to another, that is, we can fix \( \varepsilon_i \) to always be standard normal and let sigma change in value to accommodate difference in variability:

\[
y_i = \mu + \beta_1 X_{1i} + \cdots + \beta_k X_{ki} + \sigma \varepsilon_i
\]

where \( \varepsilon_i \sim N(0,1) \) is a random error term and \( \sigma > 0 \) is a scale parameter. However, since survival data cannot be negative, it is common to use \( y_i = \log(t_i) \), for example, as the dependent variable. This type of survival model linearized by taking logs on survival time
to ensure positive predictive values are AFT models, and the covariates act in a multiplicative manner on time. The distribution of T-RTP must be determined (e.g., Weibull) to fit the appropriate AFT model. In contrast with Cox models that provide estimates of the “hazard” associated with RTP, AFT models provide an estimate of the expected time to return to play and may be more robust if the accurate T-RTP distribution is specified.

5.4.2 Sports Medicine Research Institute

While the analyses and findings in this dissertation are limited to the high school athlete population, these T-RTP analysis methods can be applied to other populations where return to activity after physical injury is of interest. The Sports Medicine Research Institute (SMRI), housed within the College of Health Sciences at the University of Kentucky, is a research center focused on injury prevention, treatment, rehabilitation, and performance for tactical, collegiate, and youth athletes [96]. The musculoskeletal and neuro-cognition research groups within the SMRI are actively participating in research studies related to return to play in athletes and return to duty in the military population. More accurate and meaningful results from time to event analyses in future studies can be obtained by considering and addressing the data limitations identified in this dissertation, particularly the convention of measuring T-RTP in discrete time intervals. The time to event analysis methods detailed in this dissertation are directly applicable for both research groups within the SMRI, and the results are translatable to better clinical treatment, stronger athlete adherence to rehabilitation protocols, and prevention of reinjury.
Table 5.1 Comparison of RTP Incidence Proportion (IP) to Lifetable RTP probabilities by ankle sprain severity

<table>
<thead>
<tr>
<th>T-RTP</th>
<th>Known T-RTP(^1) (n)</th>
<th>Censored RTP (n)</th>
<th>Total “at risk” for RTP(^2) (n)</th>
<th>Estimated IP of RTP(^3) (95% CI)</th>
<th>Estimated Risk of RTP (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>(1^{\text{st}}) degree LAS (N=1305)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1-2 days</td>
<td>221</td>
<td>1</td>
<td>1305</td>
<td>.17 (.15, .19)</td>
<td>.17 (.15, .19)</td>
</tr>
<tr>
<td>3-6 days</td>
<td>450</td>
<td>4</td>
<td>1084</td>
<td>.42 (.39, .44)</td>
<td>.42 (.39, .45)</td>
</tr>
<tr>
<td>7-9 days</td>
<td>268</td>
<td>1</td>
<td>634</td>
<td>.42 (.38, .46)</td>
<td>.43 (.39, .47)</td>
</tr>
<tr>
<td>10-21 days</td>
<td>252</td>
<td>10</td>
<td>366</td>
<td>.69 (.64, .74)</td>
<td>.71 (.66, .76)</td>
</tr>
<tr>
<td>&gt;21 days</td>
<td>38</td>
<td>60</td>
<td>114</td>
<td>.33 (.25, .42)</td>
<td>.56 (.44, .68)</td>
</tr>
<tr>
<td>(2^{\text{nd}}) degree LAS (N=645)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1-2 days</td>
<td>66</td>
<td>1</td>
<td>645</td>
<td>.10 (.08, .13)</td>
<td>.10 (.08, .13)</td>
</tr>
<tr>
<td>3-6 days</td>
<td>175</td>
<td>2</td>
<td>579</td>
<td>.30 (.26, .34)</td>
<td>.30 (.27, .34)</td>
</tr>
<tr>
<td>7-9 days</td>
<td>128</td>
<td>0</td>
<td>404</td>
<td>.32 (.27, .36)</td>
<td>.32 (.27, .36)</td>
</tr>
<tr>
<td>10-21 days</td>
<td>196</td>
<td>3</td>
<td>276</td>
<td>.71 (.66, .76)</td>
<td>.72 (.67, .78)</td>
</tr>
<tr>
<td>&gt;21 days</td>
<td>42</td>
<td>32</td>
<td>80</td>
<td>.53 (.42, .63)</td>
<td>.72 (.61, .84)</td>
</tr>
<tr>
<td>(3^{\text{rd}}) degree LAS (N=136)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1-2 days</td>
<td>4</td>
<td>0</td>
<td>136</td>
<td>.03 (.00, .06)</td>
<td>.03 (.00, .06)</td>
</tr>
<tr>
<td>3-6 days</td>
<td>27</td>
<td>1</td>
<td>132</td>
<td>.20 (.14, .27)</td>
<td>.21 (.14, .27)</td>
</tr>
<tr>
<td>7-9 days</td>
<td>24</td>
<td>0</td>
<td>105</td>
<td>.23 (.15, .31)</td>
<td>.23 (.15, .31)</td>
</tr>
<tr>
<td>10-21 days</td>
<td>42</td>
<td>1</td>
<td>81</td>
<td>.52 (.41, .63)</td>
<td>.53 (.42, .64)</td>
</tr>
<tr>
<td>&gt;21 days</td>
<td>24</td>
<td>13</td>
<td>39</td>
<td>.62 (.46, .77)</td>
<td>.79 (.64, .93)</td>
</tr>
</tbody>
</table>

\(^1\)Number of injured athletes in which RTP date is known at beginning of time interval.

\(^2\)Number of injured athletes with no RTP with known date at beginning of time interval.

\(^3\)Estimated epidemiological incidence proportion for RTP [27].

RTP=Return to Play
References


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**Regional**

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**Morris, S.** (November 29, 2016) “An Introduction to Power Analysis and Sample Size Calculations using nQuery.” University of Kentucky Center for Health Services Research Data, Analytics, and Statistical Core Deep Dive Presentation Series, Lexington, KY.

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Poster Presentations
Regional
National


